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MEASURING AND EXPLAINING MANAGEMENT PRACTICES ACROSS FIRMS AND COUNTRIES*

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We use an innovative survey tool to collect *management practice* data from 732 medium-sized firms in the United States, France, Germany, and the United Kingdom. These measures of managerial practice are strongly associated with firm-level productivity, profitability, Tobin's Q , and survival rates. Management practices also display significant cross-country differences, with U.S. firms on average better managed than European firms, and significant within-country differences, with a long tail of extremely badly managed firms. We find that poor management practices are more prevalent when product market competition is weak and/or when family-owned firms pass management control down to the eldest sons (primogeniture).

I. INTRODUCTION

Economists have long speculated on why such astounding differences in productivity performance exist between firms and plants within countries, even within narrowly defined sectors. For example, labor productivity varies dramatically even within the

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same five-digit industry, and these differences are often highly persistent over time.¹

The focus of much applied economic research has been in “chipping away” at these productivity differences through better measures of inputs (capital, materials, skills, etc.). Some parts of the literature have attempted to see how much of the residual can be accounted for by explicit measures of technology, such as research and development or information and communication technologies. But technology is only one part of the story, and a substantial unexplained productivity differential still remains, which panel data econometricians often label as the fixed effects of “managerial quality” (e.g., Mundlak [1961]).

While the popular press and business schools have long stressed the importance of good management, empirical economists have had relatively little to say about management practices. A major problem has been the absence of high-quality data that are measured in a consistent way across countries and firms. One of the purposes of this paper is to present a survey instrument for the measurement of managerial practices. We collect original data using this survey instrument from a sample of 732 medium-sized manufacturing firms in the United States, the United Kingdom, France, and Germany.

We start by evaluating the quality of these survey data. We first conduct internal validation by resurveying firms to interview different managers in different plants using different interviewers in the same firms and find a strong correlation between these two independently collected measures. We then conduct external validation by matching the survey data with information on firm accounts and stock market values to investigate the association between our measure of managerial practices and firm performance. We find that better management practices are significantly associated with higher productivity, profitability, Tobin’s Q , sales growth rates, and firm-survival rates. This is true in both our English-speaking countries (the United Kingdom and the United States) and the continental European countries (France and Germany), which suggests that our characterization of good management is not specific to Anglo-Saxon cultures.

We then turn to analyzing the raw survey data and observe a surprisingly large spread in management practices across firms (see Figure I). Most notably, we see a large number of firms that

1. For example, Baily, Hulten, and Campbell (1992), Bartelsman and Dhrymes (1998), Bartelsman and Doms (2000), Foster, Haltiwanger, and Syverson (2005).

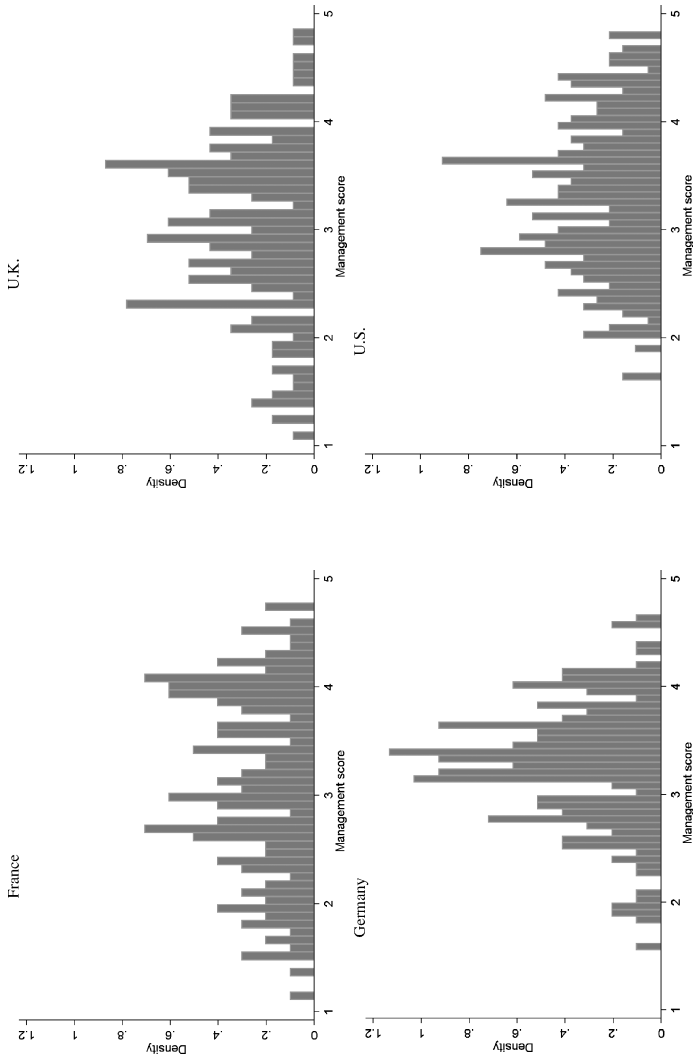


FIGURE I
Distribution of Management Scores by Country

Notes: These are the distributions of the raw management scores (simple averages across all 18 practices for each firm). 1 indicates worst practice, 5 indicates best practice. There are 135 French observations, 156 German observations, 151 UK observations, and 290 U.S. observations.

appear to be extremely badly managed, with ineffective monitoring, targets and incentives. We also observe significant variations in management practices across our sample of countries, with U.S. firms on average better managed than European firms.

This raises the main question that we address in the paper—what could rationalize such variations in management practices? The two factors that appear to play an important role are product market competition and family firms. First, higher levels of competition (measured using a variety of different proxies, such as trade openness) are strongly associated with better management practices. This competition effect could arise through a number of channels, including the more rapid exit of badly managed firms and/or the inducement of greater managerial effort. Second, family-owned firms in which the chief executive officer (CEO) is chosen by *primogeniture* (the eldest male child) tend to be very badly managed. In theory, family *ownership* could have beneficial effects from the concentration of ownership, as this may overcome some of the principal-agent problems associated with dispersed ownership. In our data, we find that family ownership combined with professional management (i.e., where the CEO is not a family member) has a mildly positive association with good managerial practices. The impact of family *ownership and management* is more theoretically ambiguous, however, with positive effects from reducing the principal-agent problem but negative effects due to more limited selection into managerial positions as well as the Carnegie effect.² Empirically, we find that companies that select the CEO from all family members are no worse managed than other firms, but those that select the CEO based on primogeniture are very poorly managed.

The impact of competition and family firms is quantitatively important. Low competition and primogeniture in family firms account for about half of the tail of poorly performing firms. Across countries, competition and family firms also play a large role, accounting for over half of the gap in management practices between the United States and France and one-third of the gap between the United States and the United Kingdom. One reason is that European competition levels are lower than those in the United States. Another reason is that primogeniture is much more common in

2. The “Carnegie effect” is named after the great philanthropist Andrew Carnegie, who claimed, “The parent who leaves his son enormous wealth generally deadens the talents and energies of the son, and tempts him to lead a less useful and less worthy life than he otherwise would.” See also Holtz-Eakin, Joulfaian, and Rosen (1993).

France and the United Kingdom due to their Norman heritage, in which primogeniture was legally enforced to preserve concentrated land-holdings for military support. More recently, Britain and other European countries have also provided generous estate tax exemptions for family firms.

Our work relates to a number of strands in the literature. First, our findings are consistent with recent econometric work looking at the importance of product market competition in increasing productivity.³ It has often been speculated that the productivity-enhancing effects of competition work through improving average management practices, and our study provides support for this view. Second, economic historians such as Landes (1969) and Chandler (1994) have claimed that the relative industrial decline of the United Kingdom and France in the early twentieth century was driven by their emphasis on family management, compared to the German and American approach of employing professional managers.⁴ Our results suggest this phenomenon is still important almost a century later. A third related strand is work on the impact of human resource management (HRM),⁵ which also finds that these management practices are linked to firm performance. Finally, there is the recent contribution of Bertrand and Schoar (2003), who focus on the impact of changing CEOs and CFOs in very large quoted U.S. firms. This will tend to reflect the impact of management *styles* and *strategies*, complementing our work emphasizing the *practices* of middle management. We see management practices as more than the attributes of the top managers: they are part of the organizational structure and behavior of the firm, typically evolving slowly over time even as CEOs and CFOs come and go.

The layout of this paper is as follows. Section II discusses why management practices could vary, Section III discusses measuring management practices with our management data, and Section IV offers an external validation of the survey tool. In Section V, we discuss the distribution of management practices and offer evidence on the causes for the variations in management. In Section VI, we pull this all together to try to explain

3. There is a very large number of papers in this area, but examples of key recent contributions would be Nickell (1996), Olley and Pakes (1996), and Syverson (2004a, 2004b).

4. See also the recent literature on family firms and performance, for example, Bertrand et al. (2005), Morck, Wolfenzon, and Yeung (2005), Perez-Gonzalez (2005), and Villalonga and Amit (2005).

5. For example, Ichinowski, Shaw, and Prenushi (1997), Lazear (2000), Black and Lynch (2001), and Bartel, Ichinowski, and Shaw (2005).

management practices across firms and countries. Finally, some concluding comments are offered in Section VII. More details of the data, models, and results can be found in the appendixes and the working paper version.

II. MODELS OF MANAGEMENT PRACTICES

II.A. Why Are There Good and Bad Management Practices?

Our starting point is that there are likely to be management practices that are, on average, “good” for firm productivity. Organizations where managers are of high quality or supply effort that is more effective will tend to have better managerial practices. This notion underlies the Lucas (1978) model of firm size and Mundlak’s (1961) discussion of firm fixed effects. It is also inherent in the benchmarking exercises that are ubiquitous in the business world. We will discuss in detail the challenge of empirically measuring these, but first consider some examples. Japanese lean manufacturing techniques (just-in-time, quality circles, etc.) were a managerial innovation that was initially resisted but gradually became adopted across the West, first in the automobile industry and then elsewhere. Eventually these managerial methods were acknowledged to be generally superior, even if they are not always adopted (we discuss reasons for this below). A second example would be performance tracking, where a firm systematically collects, analyzes, and communicates key performance indicators (KPIs). The absence of any easily collected and analytically useful measures of firm performance is likely to indicate poor management. A third example is promotion decisions. Promoting workers who are poor performers or simply because of their tenure in the firm is likely to lead to lower productivity than considering individual performance when deciding whether to move an employee up the hierarchy.

If certain management practices are beneficial for productivity, why do all firms not immediately adopt them? There are static and dynamic reasons for this. On the static side, there are at least three reasons that an industry will not adopt best managerial practices, even in the long run—costs, agency considerations, and industry heterogeneity. First, although a management practice may be beneficial for productivity, there are also costs to take into account. Upgrading management is a costly investment and some firms may simply find that these costs outweigh the benefits of moving to better practices. In other words, although

improving management practices increase productivity, profits will not rise.

Second, bad practices may be attractive to managers because of the greater effort involved in moving to best practice. Agency considerations can drive a wedge between shareholder interest and management behavior, and there may not be a contractual solution to obtain optimal managerial effort to improve practices. A large literature discusses the theoretical and empirical importance of managerial entrenchment, and we discuss why low product market competition and the prevalence of family firms may make firm value maximization less likely.

A third reason that firms may not adopt best practice is simple heterogeneity. The optimal level of practices may vary due to differential costs and/or benefits. For example, investing heavily in best practice “people management” through rigorous appraisals will be less beneficial if workers are unskilled and quite homogeneous. In the results section we examine this idea by looking at how different types of people management practices vary systematically with skill intensity in the environment.

In a dynamic context, frictions will slow down the adoption of best management practice. Even if a new management practice were a purely technological innovation, we would expect it to take time to spread throughout the economy (recall the lean manufacturing example). First, there may be learning effects, as information about the new management practice diffuses only slowly across firms. Second, there are costs of adjustment that will mean that moving immediately to the best practice is unlikely to be optimal. One extreme form of adjustment costs is when only new entrants are able to implement the best practice, as incumbent firms keep to the same practices that were imprinted upon them by their founding entrepreneurs (cf. Jovanovic [1982]). In this case, a selection mechanism will gradually allocate more production to the new firms with better practices and away from the incumbents (e.g., Hoppenhayn [1992]). Selection is likely to be an important way in which management practices spread, even in models where incumbents can learn to improve, as the learning process will still take time.

II.B. The Determinants of Management Practices: Competition and Family Firms

We focus on product market competition and family firms as reasons for the distribution of management practices across firms and countries, as these have been the subject of much theoretical

discussion and are important in our data. We investigated a large number of other possible factors that we discuss in the results section (e.g., corporate governance, labor unions, capital markets, and job regulations). These appeared to be empirically less important in the data than competition and family firms. This may be because the effects of these other factors are more subtle, and given our current sample size, we are not able to statistically identify their effects. In 2006 we conducted a second wave of the survey, increasing the sample size almost fivefold, that will, we hope, enable a more detailed future investigation of alternative influences on management practices.

Product Market Competition. The most obvious effect of competition on management is through a Darwinian selection process, as discussed in the dynamic “frictions” model of selection. Higher product market competition will drive inefficient firms out of the market and allocate greater market share to the more efficient firms. Syverson (2004a, 2004b) focuses on productivity and offers supportive evidence for these predictions in his analysis of the U.S. cement industry, finding that tougher competition is associated with both a higher average level of productivity and a lower dispersion of productivity, as the less efficient tail of firms have been selected out. Therefore, we expect a better average level (and a more compressed spread) of management practices in environments that are more competitive.

Competition could also affect the degree of managerial effort under agency cost models, although formally its impact is ambiguous. Higher competition can increase managerial effort, as the fear of bankruptcy is higher (Schmidt 1997). In addition, the sensitivity of market share to marginal cost differences is greater under higher competition, so this increases the marginal return to managerial effort. On the other hand, profit margins will be lower when competition is more intense, so the rewards of the profit-related component of pay will also be lower, and this will tend to depress managerial effort. Because of these offsetting influences, the effect of competition on effort cannot in general be signed. Recent contributions that allow for endogenous entry, however, tend to find that the pro-effort effect will dominate when within-market competition increases (say, from a fall in transport costs). This is because the fall in margins will mean that in equilibrium, firm size will increase, so a unit decrease in marginal costs through greater managerial effort is more valuable (e.g., Raith [2003]; Vives [2005]; Bloom and Van Reenen [2006]).

Family Firms. The theoretical implications of family ownership depend on the extent of involvement in management. Family ownership *per se* may have advantages over dispersed ownership because the (concentrated) ownership structure may lead to closer monitoring of managers (e.g., Berle and Means [1932]).⁶ Furthermore, under imperfect capital markets, founders will find it difficult to sell off the firm to outside investors (Caselli and Gennaioli 2006). Moreover, when minority investor rights are not well protected, it may be difficult to diversify ownership, so family firms may be optimal in a second-best world (Burkart, Panunzi, and Shleifer 2003).

Even when a firm is family-owned, outside professional managers can be appointed to run the firm, as is common in Germany, for example (see Section V.C). Combining family ownership with family management has several potential costs. Selecting managers only from family members limits the pool of potential talent to run the firm, and there is less competition for senior positions. Furthermore, the knowledge that family members will receive management positions in the future may generate a “Carnegie effect” of reducing their investment in human capital earlier in life. These selection and Carnegie effects are likely to be much more negative for primogeniture family firms, in which the eldest son is destined to control the firm from birth. On the other hand, principal-agent problems may be mitigated by combining ownership and control (e.g., in the model of Burkart, Panunzi, and Shleifer [2003]). There may also be investment in firm-specific human capital if the owners’ children expect to inherit the family firm. So ultimately, the impact of family firms on management practices is an empirical matter.

Family-owned firms should have incentives to balance these factors optimally before deciding on using family or external managers. However, companies may choose family management even though this is suboptimal for company performance because family members receive amenity value from managing the family firm, which often bears the family name and has been managed by several previous generations. In this case, the family may accept lower economic returns from their capital in return for the

6. Bennedsen et al. (2007) list a range of additional potential benefits (and costs) of family ownership, although these are likely to be less important than those discussed in the main text. The benefits include working harder due to higher levels of shame from failure, trust and loyalty of key stakeholders, and business knowledge from having grown up close to the firm. The costs include potential conflicts between business norms and family traditions.

private utility of managerial control. Indeed, the desire to retain family management may also be a reason for the refusal of family owners to sell equity stakes in the company to outsiders.

The existing evidence on inherited family firms suggests that family *ownership* has a mixed effect on firm profitability, but family *management* appears to have a substantially negative effect.⁷ Our approach in this paper is to examine the impact of family firms on management practices directly rather than only look at firm performance measures. Although there may be some endogeneity problems with the family-firms effect on management, these selection effects seem to cause OLS estimates to *underestimate* the damage of family involvement in management. This is because family firms are empirically more likely to involve professional managers when the firm has suffered a negative shock (see Bennedsen et al. [2007]).⁸

III. MEASURING MANAGEMENT PRACTICES

To investigate these issues, we first have to construct a robust measure of management practices that overcomes three hurdles: scoring management practices, collecting accurate responses, and obtaining interviews with managers. We discuss these issues in turn.

III.A. Scoring Management Practices

To measure management requires codifying the concept of “good” or “bad” management into a measure applicable to different firms across the manufacturing sector. This is a hard task, as good management is tough to define and is often contingent on a firm’s environment. Our initial hypothesis was that while some management practices are too contingent to be evaluated as “good” or “bad,” others can potentially be defined in these terms, and it is these practices we tried to focus on in the survey. To do this we used a practice evaluation tool developed by a leading international management consultancy firm. In order to prevent any perception of bias with our study we chose to receive no financial support from this firm.

7. See for example Perez-Gonzalez (2005) and Villalonga and Amit (2005).

8. Bennedsen et al. (2007) construct a dataset of more than 6,000 Danish firms, including information on the gender of the first-born child, which they use as an instrumental variable for remaining under family management after a succession.

The practice evaluation tool defines and scores from one (worst practice) to five (best practice) across eighteen key management practices used by industrial firms. In Appendix I.A we detail the practices and the type of questions we asked in the same order as they appeared in the survey. In Appendix I.B we give four example practices, the associated questions and scoring system, and three anonymized responses per practice. Bloom and Van Reenen (2006) give examples that are more extensive across all eighteen practices.

These practices are grouped into four areas: *operations* (three practices), *monitoring* (five practices), *targets* (five practices), and *incentives* (five practices). The shop-floor operations section focuses on the introduction of lean manufacturing techniques, the documentation of processes improvements, and the rationale behind introductions of improvements. The monitoring section focuses on the tracking of performance of individuals, reviewing performance (e.g., through regular appraisals and job plans), and consequence management (e.g., making sure that plans are kept and appropriate sanctions and rewards are in place). The targets section examines the type of targets (whether goals are simply financial or operational or more holistic), the realism of the targets (stretching, unrealistic, or nonbinding), the transparency of targets (simple or complex), and the range and interconnection of targets (e.g., whether they are given consistently throughout the organization). Finally, the incentives section includes promotion criteria (e.g., purely tenure-based or including an element linked to individual performance), pay and bonuses, and fixing or firing bad performers, where best practice is deemed the approach that gives strong rewards to those with both ability and effort. A subset of the practices has similarities to those used in studies on human resource management practices.

Since the scaling may vary across practices in the econometric estimation, we convert the scores (from the one to five scale) to z -scores by normalizing by practice to mean zero and standard deviation one. In our main econometric specifications, we take the unweighted average across all z -scores as our primary measure of overall managerial practice, but we also experiment with other weighting schemes based on factor analytic approaches.

There is scope for legitimate disagreement over whether all of these measures really constitute “good practice.” Therefore, an important way to examine the external validity of the measures is to examine whether they are correlated with data on firm performance constructed from completely independent data

sources—company accounts and the stock market. We do this in Section IV.

III.B. Collecting Accurate Responses

With this evaluation tool, we can provide some quantification of firms' management practices. However, an important issue is the extent to which we can obtain unbiased responses from firms to our questions. In particular, will respondents provide accurate responses? As is well known in the surveying literature (e.g., Bertrand and Mullainathan [2001]), a respondent's answer to survey questions is typically biased by the scoring grid, anchored toward those answers that the respondent expects the interviewer to think are correct. In addition, interviewers may themselves have preconceptions about the performance of the firms they are interviewing and bias their scores based on their *ex ante* perceptions. More generally, a range of background characteristics, potentially correlated with good and bad managers, may generate some kinds of systematic bias in the survey data.

To try to address these issues, we took a range of steps to obtain accurate data. First, the survey was conducted by telephone, without telling the managers they were being scored.⁹ This enabled scoring to be based on the interviewer's evaluation of the firm's actual practices, rather than its aspirations, the manager's perceptions, or the interviewer's impressions. To run this "blind" scoring we used open questions (e.g., "can you tell me how you promote your employees?") rather than closed questions (e.g., "do you promote your employees on tenure [yes/no]?"). Furthermore, these questions target actual practices and examples, with the discussion continuing until the interviewer can make an accurate assessment of the firm's typical practices based on these examples. For each practice, the first question is broad, with detailed follow-up questions to fine-tune the scoring. For example, in dimension (1), modern manufacturing introduction, the initial question is "Can you tell me about your manufacturing process?" and is followed up by questions such as "How do you manage your inventory levels?"

Second, the interviewers did not know anything about the firm's financial information or performance in advance of the

9. This survey tool has been passed by Stanford's Human Subjects Committee. The deception involved was deemed acceptable because it (i) is necessary to get unbiased responses; (ii) is minimized to the management practice questions and temporary (we send managers debriefing packs afterward); and (iii) presents no risk, as the data are confidential.

interview. This was achieved by selecting medium-sized manufacturing firms and by providing only firm names and contact details to the interviewers (but no financial details). Consequently, the survey tool is “double blind”—managers do not know they are being scored and interviewers do not know the performance of the firm. The interviewers were incentivized on the number of interviews they ran and so had no interest in spending time researching the companies in advance of running the interview. These medium-sized firms (the median size was 675 employees) would not be known by name and are rarely reported in the business media. The interviewers were specially trained graduate students from top European and U.S. business schools. All interviews were conducted in the manager’s native language.

Third, each interviewer ran over 50 interviews on average, allowing us to remove interviewer fixed effects from all empirical specifications. This helps to address concerns over inconsistent interpretation of categorical responses (see Manski [2004]), standardizing the scoring system.

Fourth, the survey instrument was targeted at plant managers, who are typically senior enough to have an overview of management practices but not so senior as to be detached from day-to-day operations of the enterprise.

Fifth, we collected a detailed set of information on the interview process itself (number and type of prior contacts before obtaining the interviews, duration, local time of day, date, and day of the week), on the manager (gender, seniority, nationality, company and job tenure, internal and external employment experience, and location), and on the interviewer (individual interviewer fixed effects, time of day, and subjective reliability score). Some of these survey controls are significantly informative about the management score¹⁰ and help reduce residual variation.

III.C. Obtaining Interviews with Managers

Each interview took on average fifty minutes and was run in the summer of 2004 from the Centre for Economic Performance at the London School of Economics. Overall, we obtained a relatively high response rate of 54%, which was achieved through four steps.

10. In particular, we found that the scores were significantly higher for senior managers when interviews were conducted later in the week and/or earlier in the day. That is to say, scores were highest, on average, for senior managers on a Friday morning and lowest for junior managers on a Monday afternoon. By including information on these characteristics in our analysis, we explicitly controlled for these types of interview bias.

First, the interview was introduced as “a piece of work”¹¹ without discussion of the firm’s financial position or its company accounts, making it relatively noncontroversial for managers to participate. Interviewers did not discuss financials in the interviews, both to maximize the participation of firms and to ensure that our interviewers were truly blind to the firm’s financial position. Second, practices were ordered to lead with the least controversial (shop-floor operations management) and finish with the most controversial (pay, promotions, and firings). Third, interviewers’ performance was monitored, as was the proportion of interviews achieved, so they were persistent in chasing firms (the median number of contacts each interviewer made in setting up the interview was 6.4). The questions are also about practices within the firm, so that any plant managers can respond, so there were potentially several managers per firm who could be contacted.¹² Fourth, the written endorsement of the Bundesbank (in Germany) and the Treasury (in the United Kingdom) and a scheduled presentation to the Banque de France helped demonstrate to managers that this was an important exercise with official support.

III.D. Sampling Frame and Additional Data

Since our aim is to compare across countries, we decided to focus on the manufacturing sector, where productivity is easier to measure than in the nonmanufacturing sector. We also focused on medium-sized firms, selecting a sample where employment ranged between 50 and 10,000 workers (with a median of 675). Very small firms have few publicly available data. Very large firms are likely to be more heterogeneous across plants, and so it would be more difficult to get a picture of managerial performance in the firm as a whole from one or two plant interviews. We drew a sampling frame from each country to be representative of medium-sized manufacturing firms and then randomly chose the order of which firms to contact (see Appendix II for details). We also excluded any clients of our partnering consultancy firm from our sampling frame. Since we used different databases in Europe (Amadeus) and the United States (Compustat), we had concerns regarding the cross-country comparisons, so we include

11. We avoided using the words “research” or “survey,” as many firms link these to market research surveys, which they usually refuse to be involved with.

12. We found no significant correlation between the number, type, and time span of contacts before an interview is conducted and the management score. This suggests that while different managers may respond differently to the interview proposition, this does not appear to be directly correlated with their responses or the average management practices of the firm.

country dummies in all regression tables. The only exception is Table VI, where we are explicitly comparing the national averages, and here (as elsewhere) we are careful to include controls for size and listing status.

In addition to the standard information on management practices, we also ran two other surveys with the same firm (details in Bloom and Van Reenen [2006]). First, we collected information from a separate telephone survey of the human resource department on the average characteristics of workers and managers in the firm, such as gender, age, college degree, hours, holidays, sickness, occupational breakdown, and a range of questions on the organizational structure of the firm and the work-life balance. Second, we collected information from public data sources and another telephone survey in summer 2005 on family ownership, management, and succession procedures, typically answered by the CEO or his office. Quantitative information on firm sales, employment, capital, materials, and so forth came from the company accounts and proxy statements, while industry level data came from the OECD. To control for industry heterogeneity, we condition on a full set of three-digit industry dummies (105 in all). As a robustness check, we also considered the subsample where we have at least five sampled firms in every three-digit industry (582 firms from our main sample of 732 firms). All of the reported results are as strong, if not stronger, for this subsample.

Comparing the responding firms with those in the sampling frame, we found no evidence that the responders were systematically different from the nonresponders on any of the performance measures. They were also statistically similar on all the other observables in our dataset. The only exception was size, where our firms were slightly larger on the average than those in the sampling frame.

III.E. Evaluating and Controlling for Measurement Error

The data potentially suffer from several types of measurement error that are likely to bias the association of firm performance with management toward zero. First, we could have measurement error in the management practice scores obtained using our survey tool. To quantify this, we performed repeat interviews on 64 firms, contacting different managers in the firm, typically at different plants, using different interviewers. To the extent that our management measure is truly picking up general company-wide management practices, these two scores should be

correlated, while to the extent that the measure is driven by noise, the measures should be independent.

The correlation of the first interviews with the second interviews was strongly positive (a correlation coefficient of .734 with a p -value of .000). Furthermore, there is no obvious (or statistically significant) relationship between the degree of measurement error and the absolute score. That is, high and low scores appear to be as well measured as average scores, and firms that have high (or low) scores on the first interview tend to have high (or low) scores on the second interview. Thus, firms that score below two or above four appear to be genuinely badly or well managed rather than extreme draws of sampling measurement error.

Analyzing the measurement error in more detail, we find that the practice level measures are noisier, with 42% of the variation in the scores due to measurement error, compared to the average firm's scores, with 25% of the variation due to measurement error. This improved the signal-noise ratio in the firm-level average measure—which is our primary management proxy—is due to the partial averaging out of measurement errors across practices.

The second type of measurement error concerns the fact that our management practices cover only a subset of all management practices that drive performance. For example, our interviews did not contain any questions on management strategy (such as pricing or merger and acquisition policies). However, so long as firms' capabilities across all management practices are positively correlated—which they are, significantly, within the eighteen practices examined—our measure based on a subset of practices will provide a proxy of the firm's true management capabilities.

IV. VALIDATING THE MANAGEMENT PRACTICE DATA

Before we investigate the reasons for the spread of management practices across firms, it is worth evaluating whether these practices are correlated with firm performance. The purpose of this exercise is not to directly identify a causal relationship between our management practice measures and firm performance. It is rather an external validity test of the survey measurement tool to check that the scores are not just "cheap talk" but are actually correlated with quantitative measures of firm performance from independent data sources on company accounts, survival rates, and market value.

IV.A. *Econometric Modeling of Productivity*

Consider the basic firm production function

$$(1) \quad y_{it}^c = \alpha_l^c l_{it}^c + \alpha_k^c k_{it}^c + \alpha_n^c n_{it}^c + \beta^c M_i^c + \gamma^c Z_{it}^c + u_{it}^c,$$

where Y = deflated sales, L = labor, K = capital, and N = intermediate inputs (materials) of firm i at time t in country c (we allow country-specific parameters on the inputs and in some experiments the management scores) and lower case letters denote natural logarithms ($y = \ln(Y)$, etc.). The Z s are a number of other controls that will affect productivity, such as workforce characteristics¹³ (the proportion of workers with a college degree, the proportion with MBAs, and the average hours worked), firm characteristics (firm age and whether the firm is publicly listed on the stock market), and a complete set of three-digit industry dummies and country dummies.

The crucial variable for us is management practices, denoted M . Our basic measure takes z -scores of each of the eighteen individual management practices and then averages over the variables to proxy M . We experimented with a number of other approaches, including using the primary factor from factor analysis and using the raw average management scores, and found very similar results.

The most straightforward approach to estimating equation (1) is to simply run OLS in the cross section (or on the panel with standard errors clustered by company) and assume that all the correlated heterogeneity is captured by the control variables. Since we have panel data, however, an alternative is to implement a two-step method where we estimate the production function in stage one, including fixed (or quasi-fixed) effects, and then calculate total factor productivity using the parameter estimates. We then project the “long-run” component of productivity on the management scores in a separate second step. This is the approach used by Black and Lynch (2001) in a similar two-step analysis of workplace practices and productivity. We estimate the production function in a variety of ways. The simplest method is within groups—that is, including a full set of firm dummies. We compared this to the Olley and Pakes (1996) estimator that allows

13. We experimented with a wide range of other workforce characteristics, such as gender, average worker age, and unionization. We only found measures of human capital to be statistically significant after controlling for firm characteristics. The data set and Stata estimation code are available online.

an unobserved firm-specific efficiency term to follow a first-order Markov process. Using the estimates of the production function parameters from Olley and Pakes, we construct the firm-specific efficiency measures and relate these in a second stage to management practices. Finally, we estimate using the “System GMM” approach (Blundell and Bond 2000) that also allows for the endogeneity of all the time-varying inputs (i.e., capital, labor, and materials).

IV.B. Econometric Results

Table I investigates the association between firm performance and management practices. Column (1) simply reports a level OLS specification including only labor, country, and time dummies as additional controls. The management score is strongly positively and significantly associated with higher labor productivity. The second column includes fixed capital and materials, and this almost halves the management coefficient. In column (3), we include our general controls of industry dummies, average hours worked, education, firm age, and listing status. This reduces the management coefficient slightly more, but it remains significant. Finally, in column (4), we include a set of interview noise controls to mitigate biases across interviewers and types of interviewees. This actually increases the management coefficient, as we would expect if we were stripping out some of the measurement error in the management score. Overall, the first four columns suggest that the average management score is positively and significantly correlated with total factor productivity.

In column (5) we present one example of a more econometrically sophisticated production function estimate, based on the two-step method discussed above, where we recover the unobserved long-run component of TFP and project this onto the management score and other covariates. We estimate the permanent component by the Olley-Pakes method. The results are as strong as those presented for the simple OLS regressions. The coefficient (standard error) on management was 0.071 (0.017) in a GMM version of column (5) of Table I and 0.080 (0.017) in a within-groups version. Whether estimated by GMM, Olley-Pakes, or within groups, management practices are always positively and significantly associated with the longer-run component of TFP.

We were concerned that the definition of good management may be biased toward an Anglo-Saxon view of the management world. Some may regard such business practices as suitable for Britain and America but less suitable for continental Europe.

TABLE I
ESTIMATES OF FIRM PERFORMANCE EQUATIONS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Estimation method	OLS	OLS	OLS	OLS	Olley-Pakes	OLS	OLS	Probit	OLS
Firms	All	All	All	All	All	All	Quoted	All	All
Dependent variable	Ln(Y) _{it}	Ln(Y) _{it}	Ln(Y) _{it}	Ln(Y) _{it}	Ln(Y) _{it}	ROCE	Ln(Tobin's av. Q)	Exit (by death)	Sales growth
	Sales	Sales	Sales	Sales	Sales	Profitability			
Management z-score	0.075 (0.024)	0.039 (0.012)	0.032 (0.011)	0.040 (0.012)	0.038 (0.015)	2.452 (0.676)	0.258 (0.072)	-0.200 (0.024)	0.019 (0.006)
Ln(L) _{it}	1.080 (0.034)	0.522 (0.036)	0.535 (0.033)	0.522 (0.032)	0.426 (0.022)	1.432 (1.712)	0.400 (0.194)	0.233 (0.043)	-0.021 (0.014)
Labor									
Ln(K) _{it}		0.186 (0.029)	0.147 (0.026)	0.147 (0.025)	0.158 (0.042)	-1.935 (1.390)	-0.680 (0.170)	-0.158 (0.056)	0.009 (0.012)
Capital									
Ln(N) _{it}		0.301 (0.037)	0.306 (0.026)	0.307 (0.025)	0.412 (0.026)	1.081 (1.025)	0.286 (0.110)	-0.084 (0.202)	0.008 (0.009)
Materials									
Country, time, and industry dummies	Yes	Yes	Yes	Yes	Yes	Yes	Yes	No	Yes
General controls	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Noise controls	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Firms	709	709	709	709	709	690	374	709	702
Observations	5,350	5,350	5,350	5,350	3,606	5,089	2,635	709	4,777

Notes: All columns estimated by OLS except column (8), which is estimated by probit maximum likelihood, and column (5), which is estimated using the Olley and Pakes (1996) technique. In all columns except (8), standard errors are in parentheses under coefficient estimates and allow for arbitrary heteroskedasticity and serial correlation (i.e., clustered by firm). In column (8), we report the p-value in square brackets below the marginal effects of each variable on the percentage increase in the probability of exit (between 2004 and 2005). The coefficients on capital, materials, and labor are allowed to be different across countries and consolidation status (United Kingdom is base). "General controls" comprise firm-level controls for ln(average hours worked), ln(firm age), a dummy for being listed, a dummy for consolidated accounts, the share of the workforce with degrees, and the share of the workforce with MBAs (except column (8), which just controls for ln(age) and listing status). "Noise controls" are 16 interviewer dummies, the seniority, gender, tenure and number of countries worked in of the manager who responded, the day of the week the interview was conducted, the time of the day the interview was conducted, the duration of the interview, and an indicator of the reliability of the information as coded by the interviewer. Data run between 1994 and 2004, except in column (8), which is a cross section. All regressions include a full set of three-digit industry dummies and four country dummies interacted with a full set of time dummies (except column (5), which has a linear time trend and country dummies, and column (8)). Column (5) uses a third-order series expansion in ln(capital) and ln(investment), and we also include a selection correction term following Olley and Pakes (1996). Standard errors are bootstrapped (clustered by firm) with 200 replications. After calculating the parameters of labor and materials (stage 1a) and capital (stage 1b), we calculate the efficiency term/TFP averaged by firm across all years. This is used as a dependent variable and regressed on the management score and the general controls (stage 2).

We empirically tested this by including interactions of the management term with country dummies—we could not reject the hypothesis that the coefficients on management were equal across countries.¹⁴

In addition to the overall management score, we looked at the role that individual practices play. Rerunning column (4) of Table I, we find that thirteen of the practice z -scores are individually significant at the 10% level or above, while five appear insignificant.¹⁵ The average practice-level point estimate is 0.023—about half the pooled average of 0.040—reflecting the higher practice-level measurement error. We also calculated the average score separately for the four groups of management practices and entered them one at a time into the production function. The point estimates (standard errors) were as follows: operations 0.031 (0.010), monitoring 0.025 (0.010), targets 0.032 (0.010), and incentives 0.035 (0.012).¹⁶

We also considered whether the management measure was simply proxying for better technology in the firm. Although technology measures such as research and development (R&D) and computer use are only available for subsamples of the dataset, we did not find that the management coefficient fell by very much in the production function when we included explicit measures of technology, as these are not strongly correlated with good management.¹⁷

The final four columns of Table I examine four other measures of firm performance. In column (6) we use an alternative performance measure, which is return on capital employed (ROCE), a

14. For example, we generated a dummy for the two continental European countries and interacted this with the management score. When this was entered as an additional variable in the column (4) specification, the coefficient was 0.047 with a standard error of 0.031.

15. This suggests that not all eighteen of the individual management practices are associated with better performance. We could of course construct a “refined” management measure by averaging over only the individually significant questions, but this becomes too close to crude data mining. Details of the regressions appear in Appendix I.C.

16. We also examined specifications with multiple questions or different groupings, but statistically the simple average was the best representation of the data. Part of the problem is that it is hard to reliably identify clusters of practices in the presence of measurement error. We show how subsets of management practices vary systematically in Section IV.C.

17. In the context of the specification in Table I, column (4), for the 181 firms where we observe PCs per employee, the management coefficient is 0.084, with a standard error of 0.040 (the coefficient on PCs was 0.046, with a standard error of 0.025). This compares to a management coefficient of 0.088 with a standard error of 0.041 on the same sample when PCs are not included. For the sample of 216 firms where we have R&D information, the coefficient on management is 0.043, with a standard error of 0.017, in the specification with R&D and 0.046, with a standard error of 0.017, in the specification without R&D.

profitability measure used by financial analysts and managers to benchmark firm performance (see Bertrand and Schoar [2003]). The significant and positive coefficient on management in the ROCE equation, which also includes the same set of controls as in column (4), confirms the basic productivity results. In column (7), we estimate a Tobin's Q specification (the ratio of the market value of the firm to its book value), which again includes the same set of controls as in the production function. We also find a significant and positive coefficient on management. In column (8), we estimate the relationship between exit in the twelve months after the survey and management practices. Over this period, eight firms went bankrupt, for which the implied marginal effects of management in the probit equation are large and statistically significant. In column (9), we estimate the relationship between the average annual growth rate of sales and management practices and again find a positive and significant coefficient on management. We also find a strong and positive correlation between firm size and management practices, which is consistent with the Lucas (1978) model.

The coefficients in the production function estimates are of quantitative as well as statistical significance. Although we cannot attribute causality to the management scores on productivity, a movement from the lower to the upper quartile of management scores between firms (0.972 points) is associated with an increase in productivity of between 3.2% (column (3)) and 7.5% (column (1)). Empirically the difference in TFP between the lower quartile and upper quartile of our firms is 32%. In a purely accounting sense, therefore, management scores explain between 10% and 23% of the interquartile range of productivity.

Overall, then, there is substantial evidence that the measures of management we use are positively and significantly associated with better firm performance. These results offer some external validation of the survey tool, implying that we are not simply measuring statistical noise.

IV.C. Contingent Management

In this subsection we present evidence that firms are choosing different "styles" of management systematically (cf. Athey and Stern [1998]). In particular, we hypothesize that firms in a high-skill environment may find good human-capital management practices relatively more important than those in a low-skill environment (cf. Caroli and Van Reenen [2001]).

First, we investigated the impact of the weighting across individual practices through factor analysis. There appeared to be one dominant factor that loaded heavily on all our practices—which could be labeled “good management”—that accounted for 48% of the variation.¹⁸ The only other notable factor, which accounted for a further 7% of the variation, could be labeled as “human capital management relative to fixed capital management;” it had a positive loading on most of the human-capital-oriented practices and a negative loading on the fixed capital/shop-floor operations type of practices. This second factor was uncorrelated with any productivity measures, although interestingly it was significantly positively correlated with our skills measures (e.g., the proportion of employees with college degrees) and the level of worker autonomy,¹⁹ suggesting a slightly different pattern of relative management practices across firms with different levels of human capital.

We examine this issue more explicitly in Table II, where we find robust evidence that firms with higher employee skills—as proxied by college degrees or average wages—have significantly better *relative* human-capital management practices. Column (1) regresses the average score of the three explicitly human-capital-focused practices (13, 17, and 18 in Appendix I.A) on the percentage of employees with a degree (in logs) and finds a large positive coefficient of .198. By comparison, column (2) runs the same regression but uses the average score of the three most fixed-capital-focused practices (1, 2, and 4) as the dependent variable. In this column we also find a significantly positive association, but with a smaller coefficient of .102. Column (3) uses the difference between the human-capital-focused and fixed-capital-focused management practices as the dependent variable and shows that this measure of the relative intensity of human-capital management practices (denoted “human capital – fixed capital management” in Table II) is significantly larger in highly skilled firms. Column (4) includes the general controls that weaken the correlation slightly, but it remains significant at the 10% level. Hence, while higher skilled firms have better overall management practices, they are particularly good at the most human-capital focused management practices. Column (5) repeats the specification of column (4) but uses

18. Reestimating the production functions of Table I column (4), we found that this “good management” factor score had a coefficient of 0.027, with a standard error of 0.009.

19. See Bloom et al. (2007) for a discussion of the organizational data collected in the survey.

TABLE II
SKILL-CONTINGENT MANAGEMENT PRACTICES

Dependent variable	(1)	(2)	(3)	(4)	(5)
	Human capital management	Fixed capital management	Human capital – fixed capital management	Human capital – fixed capital management	Human capital – fixed capital management
Ln(proportion of employees with college degrees)	0.198 (0.043)	0.102 (0.047)	0.096 (0.049)	0.099 (0.057)	
Ln(firm average wages) _{it}					0.340 (0.168)
General controls	No	No	No	Yes	Yes
Industry controls	No	No	No	Yes	Yes
Firms/industries	732	732	732	732	424

Notes. All columns estimated by OLS with robust standard errors in parentheses. A single cross section of data is used. "Human capital management" is the average z-score of the three explicitly human-capital-focused practices (practices 13, 17, and 18 in Appendix I.A). "Fixed capital management" is the average z-score of the three most fixed-capital-focused practices (1, 2, and 4 in Appendix I.A). "Human capital – fixed capital management" is the difference of these two averages. "General controls" comprises controls for ln(firm age), ln(average number of employees), a dummy for being listed, and a set of country dummies. "Industry controls" are a full set of three-digit industry dummies.

average wages as an alternative measure of skill. We find a similar pattern of more human-capital-focused management practices in firms with higher average wages.²⁰ Overall, Table II is consistent with a model of management practices in which firms tailor their practices to their environments.

IV.D. Firm-Performance-Related Measurement Bias

A criticism of our external validity test of looking at production functions is that for psychological reasons managers will respond optimistically in firms that are doing well even if the true state of management practices is poor. We label this phenomenon “firm-performance-related measurement bias.”

There are several considerations mitigating the problem of firm-performance-related measurement bias in our study. First, the survey is deliberately designed to try to minimize this kind of bias by using a double-blind methodology based on open questions using actual practices and examples to score the firm. So to the extent that managers talk about actual practices in their firms, this should help to reduce this measurement bias.

Second, psychological evidence (e.g., Schwarz and Strack [1999]) suggests that recent improvements in a subject’s condition are more likely to have an impact on survey responses than the absolute level of a subject’s condition. Therefore, if there were a large performance-related bias in the management scores, we would expect this to show up in recent improvements in firm productivity (relative to comparators) having a big impact on managerial responses. In fact, when we regress management scores against lagged productivity *growth* rates, there is no significant correlation. For example, a regression of management scores against the lagged productivity growth rates over the previous year generated a coefficient (standard error) of 0.108 (0.150).²¹

20. We also used a three-digit industry-level measure of skills instead of a firm-specific measure, the proportion of employees with a college degree in the United States based on data from the Current Population Survey. We found that this was also positively correlated with the relative intensity of human-capital management practices.

21. We also tested this management and productivity growth relationship over longer periods in a Table I, column (4) specification and found equally non-significant results. For example, when using the average of productivity growth in the last three years, we obtained a coefficient of 0.092 with a standard error of 0.197. The positive correlation of management with productivity levels and sales growth, but not with productivity growth, is consistent with a simple dynamic selection model. In such a model, management (and therefore productivity levels) is fixed over time, and the market gradually allocates more sales to the more productive firms.

Third, as we shall show below in Section V.B, firms in more competitive industries—defined in terms of lower historical average price-cost margins—are on average *better* managed. Therefore, at the industry level the correlation between management practices and historical average profitability goes in the reverse direction to that implied by this measurement bias story.

Finally, the appendixes in Bloom and Van Reenen (2006) report a further battery of robustness tests on this issue. For example, not all individual practices are significantly correlated with performance, as shown in the final column of Appendix I.C. Therefore, to the extent that this bias is a serious phenomenon, it only seems to affect certain practices.

In conclusion, while there is undoubtedly scope for firm-performance-related measurement bias in the survey; we do not find evidence that this is a major problem in our results.

IV.E. Reverse Causality between Management Practices and Firm Performance

Recall that it was *not* possible to regard the coefficient on management in Table I as a causal effect of management on firm performance. Our estimated effects of the “true effect” of management on productivity could be biased upward or downward due to reverse causality. For example, positive feedback could occur if higher productivity enabled cash-constrained firms to invest more resources in improving managerial practices. This would bias our coefficient on management upward. Negative feedback could occur if higher performance generated free cash flow, enabling managers to reduce their input of effort.²² This would bias the coefficient on management downward. We investigated, using product market competition and family ownership as instrumental variables for management practices (see Bloom and Van Reenen [2006] for more details). For this to be valid we need to assume that the mechanism by which competition and primogeniture family management impact on productivity is solely through improving managerial practices. Based on these admittedly very strong identification assumptions, we found that instrumental variable estimates of management were still significant at the 5% level and much larger in magnitude than the OLS coefficients (0.216 under I.V., compared to 0.042 under OLS).

22. Higher scoring practices involve more time and effort from managers on a range of monitoring and target practices, plus potentially more difficult decisions in incentive practices over hiring, firing, pay, and promotions.

V. ACCOUNTING FOR THE DISTRIBUTION OF MANAGEMENT PRACTICES

V.A. *The Distribution of Management Practices*

Having confirmed that our management measures are informative, we now proceed to examine the management scores directly. Figure I shows the distribution of the average management scores per firm across all eighteen practices, plotted by country in raw form (not in z -score form). It is clear that there is a huge amount of heterogeneity within each country, with firms spread across most of the distribution. About 2% of the overall variation in firms' average management scores is across countries, 42% is across countries by three-digit industry, and the remaining 56% is within country and industry. This spread is particularly wide when considered against the fact that a score of one indicates industry worst practice and five industry best practice. Therefore, for example, firms scoring two or less have only basic shop-floor management, very limited monitoring of processes or people, ineffective and inappropriate targets, and poor incentives and firing mechanisms. Thus, one of the central questions we address in the next section is how these firms survive.

Looking across countries, the United States has on average the highest scores (3.32), Germany is second (3.27), France third (3.11), and the United Kingdom last (3.04), with the gaps between the United States, continental Europe (France and Germany), and the United Kingdom statistically significant at the 5% level. The UK-U.S. gap also appears persistent over time. The Marshall Plan productivity mission of 1947 reported that

efficient management was the most significant factor in the American advantage [over the United Kingdom].

(Dunning 1958, p. 120)

We were concerned that some of the apparent cross-country differences in management scores might simply be driven by differences in the sampling size distribution, but these figures are robust to controls for size and whether the firm is publicly listed (see Section V.B).

The presence of the United States at the top of the ranking is consistent with anecdotal evidence from other surveys.²³ It also reflects the labor productivity rankings from other studies

23. For example, Proudfoot Consulting (2003) regularly reports that U.S. firms were least hindered by poor management practices (36%) compared to firms in Australia, France, Germany, Spain, South Africa and the United Kingdom.

comparing the four nations (the United States is at the top and the United Kingdom at the bottom). One might suspect that this was due to an “Anglo-Saxon” bias—that is why, in the previous section, we had to confront the scores with data on productivity to show that the management scores are correlated with real outcomes within countries (see Table I). Furthermore, the position of the United Kingdom as the country with the lowest average management scores indicates that the survey instrument is not intrinsically Anglo-Saxon-biased. Appendix I.C provides more details behind these cross-country comparisons and reveals a *relative* U.S. and UK strength in targets and incentives versus a German and French strength in shop-floor operations and monitoring.

V.B. Management Practices and Product Market Competition

A common argument is that variations in management practice result from the differences in product market competition, because of selection effects and/or because of variations in the incentives to supply effort. Table III attempts to investigate this by examining the relationship between product market competition and management. We use three broad measures of competition, following Nickell (1996) and Aghion et al. (2005). The first measure is the degree of import penetration, measured as the share of total imports relative to domestic production (specific to the country and the industry in which the firm operates). This is constructed for the period 1995–1999 to remove any potential contemporaneous feedback.²⁴ The second is the Lerner index of competition, which is $(1 - \text{profits/sales})$, calculated as the average across the entire firm population (excluding each firm itself). Again, this is constructed for the period 1995–1999 and is specific to the firm’s country and three-digit industry. The third measure of competition is the survey question on the number of competitors a firm faces, valued zero for “no competitors,”

Unfortunately, these samples are drawn only from the consulting group’s clients, so they suffer from serious selection bias.

24. This is measured at the ISIC-2 level, which is slightly more disaggregated than the U.S. SIC two-digit level. Melitz (2003) and others have suggested that trade exposure should truncate the lower part of the productivity distribution. We have also looked at $(\text{Imports} + \text{Exports})/\text{Production}$ as an alternative indicator of trade exposure, with results similar to those reported here.

TABLE III
MANAGEMENT AND PRODUCT MARKET COMPETITION

Estimation method Dependent variable	(1)		(2)		(3)		(4)		(5)		(6)		(7)		(8)	
	OLS z-score	Management z-score	OLS z-score	Management z-score	OLS z-score	Management z-score	OLS z-score	Management z-score	OLS z-score	Management z-score	OLS z-score	Management z-score	OLS z-score	Management z-score	OLS z-score	Management z-score
Import penetration (5-year lagged)	0.144 (0.045)	0.166 (0.071)														
Lerner index (5-year lagged)			1.516 (0.694)		1.192 (0.568)											
Number of competitors							0.143 (0.051)		0.140 (0.040)							
Firms	732	732	732	732	732	732	732	732	732	732	732	732	732	732	732	732
General controls	No	Yes	No	Yes	Yes	No	No	Yes	Yes	No	Yes	Yes	No	No	Yes	Yes

Notes: Coefficients from OLS regressions with standard errors in parentheses (robust to arbitrary heteroscedasticity and clustered by country × industry pair). Sample is a single cross section. General controls includes a full set of three-digit industry dummies, four country dummies, ln(firm size), ln(firm age), a dummy for being listed, the share of workforce with degrees, the share of workforce with MBAs, a dummy for being consolidated, and the noise controls (16 interviewer dummies, the seniority, gender, tenure, and number of countries worked in of the manager who responded, the day of the week the interview was conducted, the time of the day the interview was conducted, the duration of the interviews, and an indicator of the reliability of the information as coded by the interviewer). Import penetration = ln(import/production) in every country × industry pair with the average over 1995–1999 used. Lerner index of competition is constructed, as in Aghion et al. (2005), as the mean of (1–profit/sales) in the entire database (excluding the firm itself) for every country–industry pair (average over 1995–1999 used). Number of competitors is constructed from the response to the survey question on number of competitors, and is coded as zero for none (1% of responses), 1 for less than 5 (51% of responses), and 2 for “5 or more” (48% of responses).

one for “less than five competitors,” and two for “five or more competitors.”²⁵

In column (1) of Table III, we see that better management scores are positively and significantly associated with greater import penetration. In column (2), we reestimate the same specification but now include a full set of controls including country and industry dummies, firm size, age, and listing status. We again find that higher lagged trade competition is significantly correlated with better management. Thus, compared to other firms in the same country and industry, and after controlling for a range of firm-level characteristics, higher import penetration is significantly associated with better management scores.²⁶ In columns (3) and (4), we run two similar specifications on the lagged Lerner index of competition as an alternative competition measure and again find a significant and positive effect. In columns (5) and (6), we run two further similar specifications, but this time using managers’ own self-reported measure of the number of competitors they face, and again we find a positive and significant association: the more rivals a firm perceives it faces, the better managed it appears to be. The final two columns include all three competition measures simultaneously. Although the statistical significance and marginal effects are typically a bit lower, the same pattern of results persists. Across all columns, the conclusion emerged that tougher product market competition is associated with significantly better management practices.

The magnitude of the competition effect on average management scores is of economic as well as statistical significance. For example, in column (6) of Table III, increasing the number of competitors from “few” to “many” is associated with a management *z*-score increase of 0.140. As we will discuss later in Section VI, this lack of competition accounts for a substantial proportion of the tail of badly performing firms and the management gap between the United States and Europe.

These are conditional correlations, of course, as we have no instrumental variable for competition. However, it is likely that any

25. This question has been used, *inter alia*, by Stewart (1990) and Nickell (1996). We obtained similar results using three separate dummies for high, low, and no competitors.

26. We also experimented with many other controls (results available on request). Union density was negatively correlated with management scores, but was insignificant. Although there was a significant negative correlation between management scores and average worker age in simple specifications, this disappeared when we controlled for firm age (older workers are more likely to be matched with older firms, and older firms on average were worse managed).

endogeneity bias will cause us to *underestimate* the importance of product market competition for management. For example, in columns (3) and (4), an exogenous positive shock that raises managerial quality in an industry is likely to *increase* profitability and therefore lower the competition measure, based on the inverse Lerner index (indeed, Table I showed a positive correlation between management and individual firm-level profitability). This will make it *harder* for us to identify any positive impact of product market competition on management.²⁷

The positive effect of competition on management practices could work through two possible mechanisms: (i) increasing management scores through greater managerial effort and/or (ii) increasing the exit rate of badly managed firms relative to well managed firms (see Section II). Using average managerial hours worked as a basic proxy for effort, we find an insignificant relationship between tougher competition and longer managerial hours.²⁸ Of course, managerial hours are an imperfect proxy for managerial effort, as managers may supply more effort by a greater “intensity” of work rather than longer hours. Still, it does suggest that the margin of impact of competition is not simply on the length of the working day or week (see also Bloom, Kretschmer, and Van Reenen [2006] for further tests). Looking at the second mechanism, we did find some weak evidence that greater product market competition was associated with a reduction in the dispersion of management practices (as suggested by Figure I and by Syverson [2004a, 2004b]). For example, if we regress the coefficient of variation of management practices (in an industry-country pair) on our competition measures, there is a negative marginal effect.²⁹ This is suggestive of a selection model, where competition drives out the worst-managed firms, but again the evidence is weak, as the competition variables were not significant at

27. Similarly, better domestic management will reduce the degree of imports and enable the firm to pull away from other competitors and therefore faces fewer rivals. This will generate a bias toward zero on all the competition indicators in Table III.

28. We reestimated the specifications of Table III, columns (2), (4), and (6), using managerial hours as the dependent variable. The coefficients (standard errors) on import penetration, the Lerner index, and the number of competitors was 0.889 (0.752), -2.903 (5.664), and 0.892 (0.545), respectively. In the three regressions, one of the competition measures (the Lerner) is “incorrectly” signed and all are insignificant at the 5% level.

29. When imports were used, the coefficient was -0.043 with a standard error of 0.031, and when the Lerner index was used, the coefficient was -13.275 with a standard error of 8.943. These are estimated at the country-industry level, and we condition on having at least five firms per cell.

TABLE IV
HEREDITARY FAMILY FIRM INVOLVEMENT BY COUNTRY

%	France	Germany	UK	U.S.
Family largest shareholder	30	32	31	10
(of which) Family largest shareholder and family CEO	19	11	23	7
(of which) Family largest shareholder, family CEO, and primogeniture	14	3	15	3
Founder largest shareholder	26	5	15	18
(of which) Founder largest shareholder and CEO	19	1	12	11
Number of firms	125	152	150	290

Notes. These mean values are taken from our sample of 717 firms. Family shareholding is combined across all family members. Family involvement is defined as second-generation family or beyond. Primogeniture is defined by a positive answer to the question “How was the management of the firm passed down: was it to the eldest son or by some other way?” Alternatives to primogeniture in frequency order are younger sons, sons-in-law, daughters, brothers, wives, and nephews. “Family largest shareholder” firms defined as those with a single family (combined across all family members, who are all second generation or beyond) as the largest shareholder; “family largest shareholder and family CEO” firms are those with additionally a family member as the CEO; “family largest shareholder, family CEO, and primogeniture” with additionally the CEO selected as the eldest male child upon succession. See Appendix II for more details on construction of the variables.

conventional levels. In short, then, in samples of this size it is difficult to identify the precise mechanism through which competition has a positive effect on management practices.

V.C. Management Practices and Family Firms

There has been much recent work on the efficiency of family firms. Family firms are the typical form of ownership and management in the developing world and much of the developed world.³⁰ As Table IV shows, family involvement is common in our sample. The largest shareholding block is a family (defined as the *second* generation or beyond from the company’s founder) in around 30% of European firms and 10% of American firms. This is similar in broad magnitude to the findings of La Porta, Lopez-de-Silanes, and Shleifer (1999), who report that about 40% of medium-sized firms were family-owned in Europe and about 10% were family-owned in the United States.³¹ Interestingly, we see in the second

30. La Porta, Lopez-de-Silanes, and Shleifer (1999) and Morck, Wolfenzon, and Yeung (2005).

31. La Porta, Lopez-de-Silanes, and Shleifer (1999) define family “ownership” as controlling 20% or more of the equity; “medium-sized” as those with common equity of just above \$500 million; and “family” as including founder-owned firms. Including founder firms in our definition would increase family ownership to about 45% in Europe and 25% in the United States, higher than their numbers, although

row that many of these firms have a family member as CEO, suggesting that families are reluctant to let professional managers run their firms. In the third row, we see that in the United Kingdom and France around two-thirds of family-owned firms choose CEOs by primogeniture (succession to the eldest son), representing around 15% of the total sample. In the United States this only occurs in about one-third of the family firms, representing 3% of all firms, and in Germany only 10% of family-owned firms have primogeniture. Consequently, only 3% of German and American firms have primogeniture in our sample, compared to 14% or 15% of French and British firms. In rows (4) and (5), we look at founder firms—those companies where the largest current shareholder is the individual who founded the firm. We see that founder firms are also common in the United Kingdom and France, as well as in the United States, although much less so in Germany.

One rationale for these differences in types of family involvement across countries is the historical tradition of feudalism, particularly in the Norman societies of the United Kingdom and France. This appears to have persisted long after the Norman kingdoms collapsed, with primogeniture obligatory under English law until the Statute of Wills of 1540 and *de facto* in France until the introduction of the Napoleonic code in the early 1800s. German traditions were based more on the Teutonic principle of *gavelkind* (equal division amongst all sons). In the United States almost all the founding fathers were the younger sons of land-owning gentry, with primogeniture abolished after the Revolution ended British rule, so that equal treatment by birth order and gender was standard by the middle of the twentieth century (Menchik 1980). A second potential rationale for these differences is the structure of estate taxation, which for a typical medium-sized firm worth \$10 million or more contains no substantial family firm exemptions in the United States, but gives about a 33%, 50%, and 100% exemption in France, Germany, and the United Kingdom, respectively.

In Table V, we investigate the relationship between firms' management scores and family firms. Column (1) starts by regressing management scores against an indicator of the family as the single largest owner (defined on total family holdings)

our medium-sized firms are smaller. The main point to note is that family firms remain common in the OECD, particularly in continental Europe.

plus the standard set of control variables. We see that family ownership per se does not seem to be associated with depressed firm performance with a positive but insignificant coefficient. In column (2), we regress management practices against an indicator of family ownership and family management (defined by the CEO being a family member) and find that the coefficient becomes more negative but again is not significantly different from zero. In column (3), we include an indicator that the firm is family-owned and family-managed with the CEO succession determined by primogeniture—the current CEO is the eldest son. For these firms we see a strongly negative and significant coefficient, suggesting that the subset of family firms that adopted primogeniture successions are substantially worse managed. In column (4), we drop the general controls and show that the family firm correlation is much stronger in the unconditional regressions. In column (5), we include all three indicators and see that it is the primogeniture family firms that are driving the negative coefficients. In fact, family ownership per se has a positive association with good management. The final column drops the founder firms from the sample so that external ownership is the omitted baseline, which makes little difference to the results. Taking Table V as a whole, it seems that the combination of family ownership and primogeniture family management significantly damages company performance.

One interpretation of this result is that being a primogeniture company directly causes inferior performance in family firms due to the selection and Carnegie effects discussed in Section II. Another interpretation is that primogeniture is an indicator of firms being more generally backward, suggesting the persistence of “old-fashioned” management techniques. While this is possible, we do nevertheless find that primogeniture family firms are significantly worse managed even after including controls for firm age, average employee age, and CEO age.³² It is also difficult to see why France and the United Kingdom should exogenously have a greater number of old-fashioned firms than Germany or the

32. Another interpretation on the poor management of family firms is that they operate less formally due to a lower return from “bureaucracy” (Novaes and Zingales 2004). The point-estimates (standard errors) for the column (3) specification for individual management components are as follows: shop-floor operations, -0.434 (0.130); monitoring, -0.389 (0.117); targets, -0.242 (0.117); and incentives, -0.274 (0.096). So while there is some evidence for this in the particularly low monitoring scores for family firms, they still score significantly badly on other management components such as shop-floor operations and incentives, which are not obviously linked to more formalized management styles.

TABLE V
MANAGEMENT AND FAMILY FIRMS

	(1)	(2)	(3)	(4)	(5)	(6)
Estimation method	OLS	OLS	OLS	OLS	OLS	OLS
Sample	All	All	All	All	All	Family and external owners
Dependent variable	Management z-score	Management z-score	Management z-score	Management z-score	Management z-score	Management z-score
Family largest shareholder	0.005 (0.063)				0.138 (0.086)	0.137 (0.090)
Family largest shareholder and family CEO		-0.105 (0.075)			-0.010 (0.113)	-0.040 (0.114)
Family largest shareholder, family CEO, and primogeniture			-0.317 (0.096)	-0.590 (0.098)	-0.410 (0.122)	-0.379 (0.128)
Firms	732	732	732	732	732	618
Country controls	Yes	Yes	Yes	Yes	Yes	Yes
General controls	Yes	Yes	Yes	No	Yes	Yes

Notes. Coefficients from OLS regressions with standard errors in parentheses (robust to arbitrary heteroscedasticity). The sample is a single cross section. In columns (1) to (5), the complete sample is used; in column (6), founder firms are dropped. "General controls" are a full set of three-digit industry dummies, ln(firm size), ln(firm age), a dummy for being listed, share of workforce with degrees, share of workforce with MBAs, a dummy for being consolidated, and the noise controls (16 interviewer dummies, the seniority, gender, tenure, and number of countries worked in of the manager who responded, the day of the week the interview was conducted, the time of day the interview was conducted, the duration of the interviews, and an indicator of the reliability of the information as coded by the interviewer).

United States (given our controls for industrial structure, firm age, and size). By contrast, the common Norman legal origin of France and the United Kingdom offers a direct historical reason for the persistence of primogeniture.

Although we treat competition and primogeniture as distinct factors, the “amenity value” explanation of family firms suggests that they should be related, with family firms more prevalent when there is less competition. We find some evidence that this is the case. For example, regressing the primogeniture dummy against the same controls in Table V, we found that the number of competitors entered the regression with a coefficient of -0.038 and a standard error of 0.020 .³³ When included simultaneously in the management equation, competition and primogeniture are individually and jointly significant (F -test of 10.67), but we found no evidence of interaction effects.

V.D. Management Scores and Management Ability

An interpretation for the variation in managerial practices across firms is that our management score proxies for the underlying ability of managers (and employees) in the firm, with well managed firms those simply containing a large fraction of high-ability managers. Under this view, our proxies of human capital (such as the proportion of employees with college degrees and the proportion with MBAs) do not control for this unobserved ability. Even under this interpretation it is, of course, interesting that lower product market competition and primogeniture increase the incidence of poor-quality managers.

However, several findings cause us to doubt that the management scores we measure are simply a cipher for employee ability. First, assuming employees are paid their marginal product, we would not expect to observe the positive correlation between good management practices and profits and Tobin’s Q discussed earlier (see Table I), as this would be priced out in the market. Second, we also find that controlling for average wages has very little effect on the size of the management coefficient in the production functions, suggesting that the management score is not simply a proxy for

33. In simple primogeniture regressions using imports or the Lerner index as a market power measure and controlling for country and industry dummies, we also found that higher competition was associated with a lower probability of primogeniture. In contrast to the number of competitors, however, these competition measures were not significant at the 5% level when the complete set of controls were included.

unobserved employee ability.³⁴ Finally, CEO pay (a proxy for top managerial ability) is not correlated with our management score once we control for firm size.³⁵ Therefore, while managerial ability may account for some of the variation in management practices across firms, this is unlikely to explain all the observed variation. Our interpretation is that managerial practices are deeply embedded in the organizational capital of the firm, and this explains the higher productivity and profitability of well-managed firms. This organizational capital is greater than the sum of the parts of abilities and skills of the current employees.

VI. EXPLAINING MANAGEMENT PRACTICES ACROSS FIRMS AND COUNTRIES: QUANTIFICATION

We turn to quantifying the role of product market competition and primogeniture family firms in accounting for management practices.

VI.A. Explaining the Tail of Badly Managed Firms

One of the interesting features of the raw data is the substantial fraction of firms that appear to have surprisingly bad management practices, with scores of two or less. These firms have only basic shop-floor management, very limited monitoring of processes or people, ineffective and inappropriate targets, and poor incentives and firing mechanisms. Interestingly, most of the differences across countries highlighted in Figure I are due to the left tail³⁶—the low UK and French average

34. When we include the $\ln(\text{average wage of the firm})$ and its interactions with country dummies in a specification identical to that of column (4) in Table I, the management coefficient is 0.049 with a standard error of 0.017. This compares to a management coefficient of 0.058 with a standard error of 0.020 without the wage terms on the same sample (we only have 430 firms for this regression, compared to the 732 in Table I, because wage data are not reported for some of the firms in the sample). The wage terms are positive and significant.

35. For example, regressing $\ln(\text{CEO pay})$ on firm size, country dummies, industry dummies, and the management score, we find the coefficient (standard error) on the management score is 0.010 (0.045). Note that although CEO pay includes bonuses, it does not include share options.

36. We ran a Kolmogorov-Smirnov test of the equality of management score distributions between the United States and Germany versus the United Kingdom and France and found that this is rejected (p -value = .002) on the whole sample. If we test the equality of this distribution for management scores above two, this is not rejected (p -value = .391). After truncating at two, the coefficients on the country dummies (standard errors) in a Table VI, column (1) specification with a U.S. baseline fall to -0.015 (0.060) for Germany, -0.012 (0.078) for France, and -0.128 (0.070) for the United Kingdom, so that the U.S.-French gap is eliminated and the U.S.-UK gap falls by more than half.

management scores are primarily due to long tails of badly managed firms.

To investigate the extent to which low competition and primogeniture family firms can account for this tail of badly run firms, we split the sample based on these measures. Figure II plots the management histogram for all firms reporting low competition³⁷ and/or primogeniture family succession, accounting for 414 firms. Panel B of Figure II plots the management histogram for the remaining high-competition and no-primogeniture succession, accounting for the remaining 308 firms. Comparing these two graphs, it is clear that the tail of badly managed firms is substantially larger in the low-competition and primogeniture sample, with 9.7% of firms scoring two or less, compared to 2.9% of firms in the high-competition no-primogeniture sample.³⁸ Given that 7.0% of all firms in the sample scored two or less, controlling for competition and primogeniture succession appears to remove over half of the tail of very badly managed firms.³⁹

VI.B. Explaining the Cross-Country Variation in Management Scores

In Table VI, we attempt to account for the variations in management practices across countries. In column (1), we regress management on dummy variables for Germany, France, and the United Kingdom (with the United States omitted as the baseline category). We find that UK and French firms are significantly worse managed than U.S. firms on average, with a gap of 0.276 and 0.202, respectively. German firms are worse managed than American firms, but not significantly so with a smaller gap

37. Defined by firms reporting “few” or “no” competitors. We use this measure to analyze cross-country competition because it is consistently measured across the sample. The Lerner index and import penetration measures may vary with accounting standards and country size respectively. In the regression results, we controlled for this with country dummies and identify from within country variation, but in this section we want to look explicitly across countries.

38. This split is also true in the U.S. and European subsamples. In the United States, 5.2% of firms score two or less in the low-competition and/or primogeniture group, while 0.6% score two or less in the high-competition non-primogeniture group. In Europe, 11.2% of firms score two or less in the low-competition and/or primogeniture group, while 5.3% score two or less in the high-competition/nonprimogeniture group.

39. Competition explains around two-thirds of this reduction in the tail. Conditioning on “many” competitors alone takes the share of firms scoring two or less from 7.0% (in the whole sample) down to 4.2%.

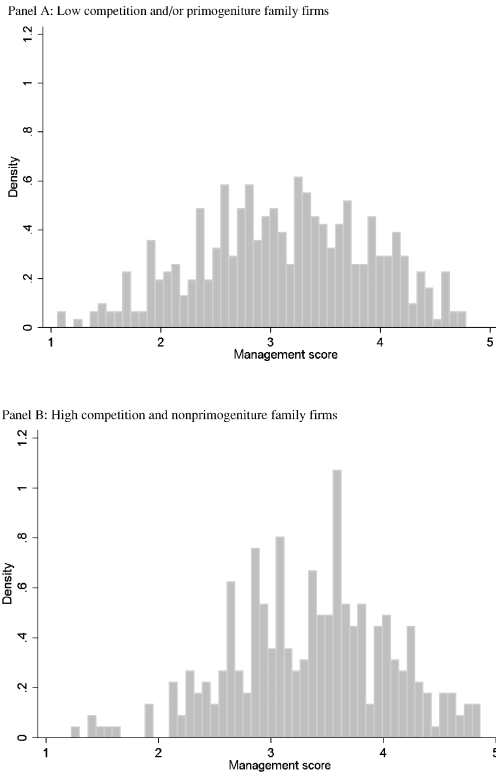


FIGURE II

The Distribution of Management Scores Split by Production Market Competition and Family Firms

Notes: Panel A shows average management scores for the 414 firms which (i) report facing “few” or “no” competitors, and/or (ii) have a family (second generation or more) as the largest shareholder with a family CEO chosen by primogeniture. Split by country is France (95), Germany (101), UK (84) and the U.S. (134). Overall 9.7% of the sample score two or less. 1 indicates worst practice, 5 indicates best practice. Panel B shows average management scores for the 308 firms which report facing “many” competitors and do not have a family (second generation or more) as the largest shareholder with a family CEO chosen by primogeniture. Split by country is France (34), Germany (51), UK (67) and the U.S. (156). Overall 2.9% of the sample score two or less. 1 indicates worst practice, 5 indicates best practice.

of 0.045. In column (2) we include controls for firm size and stock market listing status, as we were concerned that the cross-country differences might be driven by the fact that there are more smaller and unlisted firms in Europe than in the United

TABLE VI
ACCOUNTING FOR MANAGEMENT PRACTICES ACROSS COUNTRIES

	(1)	(2)	(3)	(4)	(5)
Estimation method	OLS	OLS	OLS	OLS	OLS
Dependent variable	Management raw score	Management raw score	Management raw score	Management raw score	Management raw score
	Baseline	Baseline	Baseline	Baseline	Baseline
Country is the United States	-0.045	-0.081	-0.096	-0.057	0.004
Country is Germany	(0.064)	(0.075)	(0.075)	(0.074)	(0.077)
Country is France	-0.202	-0.183	-0.136	-0.078	-0.033
	(0.086)	(0.104)	(0.104)	(0.103)	(0.103)
Country is the United Kingdom	-0.276	-0.276	-0.227	-0.196	-0.123
	(0.078)	(0.093)	(0.091)	(0.091)	(0.093)
Family largest shareholder, family CEO, and primogeniture			-0.637	-0.627	-0.582
Number of competitors			(0.101)	(0.100)	(0.098)
				0.149	0.158
				(0.052)	(0.051)
Ln(proportion of employees with degrees)					0.146
Controls for size and listed status	No	Yes	Yes	Yes	Yes
Firms	732	732	732	732	732

Notes. Coefficients from OLS regressions with standard errors in parentheses (robust to arbitrary heteroscedasticity). The sample is a single cross section. "Family largest shareholder, family CEO, and primogeniture" is a binary indicator for whether the family is the largest shareholder and the CEO is chosen by primogeniture. "Number of competitors" is constructed from the response to the survey question on number of competitors and is coded as zero for "none" (1% of responses), 1 for "less than 5" (51% of responses), and 2 for "5 or more" (48% of responses).

States. The similarity of the results to those in column (1) suggests that this is not the case.⁴⁰

In column (3), we repeat the specification of column (2) but also include a dummy for a primogeniture family firm, whose coefficient is negative and significant at the 5% level, as expected. The coefficients on the UK and French dummy variables drop substantially, by around 0.05, reflecting the extensive presence of family firms with traditional primogeniture progression rules. In column (4), we condition on our measure of the number of competitors faced by the firm. Consistent with the earlier results of the competition variables, this enters the regression with a positive and significant coefficient. The coefficient on the UK dummy drops slightly, as the degree of competition is only marginally lower in the United Kingdom than in the United States. By contrast, the coefficients for France and Germany drop by more, because the level of competition is reported to be lower by French or German companies than by U.S. firms. Together, competition and family firm status accounts for 57% ($=100 \times (0.183 - 0.078)/0.183$) of the gap between the United States and France and 29% ($=100 \times (0.276 - 0.196)/0.276$) of the gap between the United States and the United Kingdom. In column (5), we add one final control, which is the proportion of employees with a college degree, and find that this accounts for much of the remaining UK and French gap with the U.S.

Although we were expecting the competition results, the role of family firms is more surprising. The finding matches up with the earlier economic history literature of Landes (1969) and Chandler (1994), who claim that hereditary family management was probably the primary reason for the industrial decline of the United Kingdom and France relative to the United States and

40. We also considered a wide range of other checks on sampling differences. If we drop all firms with more than 1,500 workers, the mean sizes of U.S. and European firms are very similar (an average of 484 workers per firm in the United States and 504 in Europe). Rerunning column (2) on this sample gives, if anything, stronger results: the United Kingdom has a score 0.41 points below the United States, France is 0.32 below the United States, and Germany 0.26 points below (all these are differences are significant at the 5% level). Running the same specification on just the listed firms also gives similar results, with the United Kingdom and France jointly significantly worse managed than American firms at the 10% level.

Germany around the early 1900s.⁴¹ For example, Landes (1967) states that

The Britain of the late 19th century basked complacently in the sunset of economic hegemony. Now it was the turn of the third generation . . . [and] the weakness of British enterprise reflected their combination of amateurism and complacency.

(p. 563)

Before the war the model [French] enterprise was family-owned and operated, security-orientated rather than risk-taking, technologically conservative and economically inefficient.

(p. 528)

The results in Table VI suggest that family firms—at least in our sample of medium-sized manufacturing firms—are still a factor in explaining cross-country management practices one hundred years later. And extrapolating from the 30% of firms under family ownership in 2004 to the majority share they would have accounted for in the early twentieth century suggests that they could have played the dominant cross-country role in that period suggested by Landes and Chandler.

VII. CONCLUSIONS

In this paper we use an innovative survey tool to collect management practice data from 732 medium-sized manufacturing firms in Europe and the United States. The methodology described here combines the econometric advantages of large sample surveys with the measurement advantages of more detailed case-study interviews. We believe that it will be a useful part of the empirical toolkit to be used by researchers interested in the internal organization of firms. Rather than simply label unobserved heterogeneity “fixed effects,” we have explicitly developed indicators of managerial best practice.

In our application, we find that these measures of better management practice are strongly associated with superior firm performance in terms of productivity, profitability, Tobin’s Q , sales growth, and survival. We also find significant variation across countries, with U.S. firms on average much better managed than European firms. There is, however, a much larger variation

41. Nicholas (1999) provides supporting evidence for the United Kingdom, showing that over this period, individuals who inherited family firms accumulated less lifetime wealth than either firm founders or professional managers.

between firms within countries, with a long tail of extremely badly managed firms. This heterogeneity is consistent with what we know from the productivity distribution between firms and plants.

Why do so many firms exist with apparently inferior management practices, and why does this vary so much across countries? We find that this is due to a combination of (i) low product market competition, which appears to allow poor management practices to persist, and (ii) family firms passing management control down by primogeniture. European firms in our sample report facing lower levels of competition than U.S. firms. France and the United Kingdom also display substantially higher levels of primogeniture, probably due to their Norman legal origin and traditions and more generous exemption from the estate taxation regime. Product market competition and family firms alone appear to account for around half of the long tail of badly managed firms and between one-half (France) and one-third (United Kingdom) of the European management gap with respect to the United States.

Our research design focuses on managerial practices from the employer perspective rather than the worker perspective. Do these “tough” management practices come at the expense of work intensification and a breakdown of reciprocity and job satisfaction in the workplace? In a companion paper (Bloom, Kretschmer, and Van Reenen 2006), we show that our overall management score is strongly positively correlated with many pro-worker features of firms, such as more generous childcare subsidies and better work-life balance indicators. Although these indicators have no association with productivity conditional on management, it suggests that workers may prefer working in well-managed firms.

A range of potential extensions to this work is in progress, including running a second survey wave. It is important to follow up these firms in order to examine the extent to which management practice evolves over time. This will enable us to examine whether competition is working simply through selection or if there is learning of better managerial techniques by incumbent firms. The methodology of quantifying management is general enough to be applied (with modifications) to other countries and other sectors, including hospitals, schools, and the retail sector. We are also developing this survey methodology to measure the organizational structure and characteristics of firms and attempting to empirically test the long line of organizational theories of the firm.

APPENDIX I: DETAILS OF THE SURVEY QUESTIONNAIRES

APPENDIX I.A

FULL LIST OF MANAGEMENT PRACTICES WITH EXAMPLES OF THE QUESTIONS ASKED

Practice	Practice number	Practice type	Example of questions we asked
Modern manufacturing, introduction	1	Operations	Can you describe the production process for me? What kinds of lean (modern) manufacturing processes have you introduced? Can you give me specific examples? How do you manage inventory levels? What is done to balance the line?
Modern manufacturing, rationale	2	Operations	Can you talk through the rationale to introduce these processes? What factors led to the adoption of these lean (modern) management practices?
Process documentation	3	Operations	How do you go about improving the manufacturing process itself? How do problems typically get exposed and fixed? Talk me through the process for a recent problem. Do the staff ever suggest process improvements?
Performance tracking	4	Monitoring	Tell me how you track production performance. What kind of key performance indicators (KPIs) would you use for performance tracking? How frequently are these measured? Who gets to see these KPI data? If I were to walk through your factory could I tell how you were doing against your KPIs?
Performance review	5	Monitoring	How do you review your KPIs? Tell me about a recent meeting. Who is involved in these meetings? Who gets to see the results of this review?

APPENDIX I.A

(CONTINUED)

Practice	Practice number	Practice type	Example of questions we asked
Performance dialogue	6	Monitoring	How are these meetings structured? Tell me about your most recent meeting. During these meeting, how many useful data do you have? How useful do you find problem solving meetings? What type of feedback occurs in these meetings?
Consequence management	7	Monitoring	What happens if there is a part of the business (or a manager) who isn't achieving agreed upon results? Can you give me a recent example? What kind of consequences would follow such an action? Are there are any parts of the business (or managers) that seem to repeatedly fail to carry out agreed actions?
Target breadth	8	Targets	What types of targets are set for the company? What are the goals for your plant? Tell me about the financial and nonfinancial goals? What do company headquarters (CHQ) or their appropriate manager emphasize to you?
Target interconnection	9	Targets	What is the motivation behind your goals? How are these goals cascaded down to the individual workers? What are the goals of the top management team (do they even know what they are?)? How are your targets linked to company performance and their goals?
Target time horizon	10	Targets	What kind of time scale are you looking at with your targets? How are long-term goals linked to short-term goals? Could you meet all your short-run goals but miss your long-run goals?

APPENDIX I.A
(CONTINUED)

Practice	Practice number	Practice type	Example of questions we asked
Targets are stretching	11	Targets	<p>How tough are your targets? Do you feel pushed by them?</p> <p>On average, how often would you say that you meet your targets?</p> <p>Are there any targets that are obviously too easy (will always be met) or too hard (will never be met)?</p> <p>Do you feel that on targets all groups receive the same degree of difficulty?</p>
Performance clarity and comparability	12	Monitoring	<p>Do some groups get easy targets?</p> <p>What are your targets (i.e., do they know them exactly)? Tell me about them in full.</p> <p>Does everyone know their targets? Does anyone complain that the targets are too complex?</p> <p>How do people know about their own performance compared to other people's performance?</p>
Managing human capital	13	Targets	<p>Do senior managers discuss attracting and developing talented people?</p> <p>Do senior managers get any rewards for bringing in and keeping talented people in the company?</p> <p>Can you tell me about the talented people you have developed within your team?</p> <p>Did you get any rewards for this?</p>
Rewarding high performance	14	Incentives	<p>How does your appraisal system work? Tell me about the most recent round?</p> <p>How does the bonus system work?</p> <p>Are there any nonfinancial rewards for top performers?</p>

APPENDIX I.A

(CONTINUED)

Practice	Practice number	Practice type	Example of questions we asked
Removing poor performers	15	Incentives	<p>If you had a worker who could not do his job what would you do? Could you give me a recent example? How long would underperformance be tolerated? Do you find any workers who lead a sort of charmed life? Do some individuals always just manage to avoid being fixed/fired?</p>
Promoting high performers	16	Incentives	<p>Can you rise up the company rapidly if you are really good? Are there any examples you can think of? What about underperformers—do they get promoted more slowly? Are there any examples you can think of? How would you identify and develop (i.e., train) your star performers? If two people both joined the company five years ago and one was much better than the other, would he/she be promoted faster?</p>
Attracting human capital	17	Incentives	<p>What makes it distinctive to work at your company as opposed to your competitors? If you were trying to sell your firm to me how would you do this (get them to try to do this)? What don't people like about working in your firm?</p>
Retaining human capital	18	Incentives	<p>If you had a star performer who wanted to leave what would the company do? Could you give me an example of a star performers being persuaded to stay after wanting to leave? Could you give me an example of a star performer who left the company without anyone trying to keep them?</p>

Note. Scoring guide provided in Appendix I.B.

APPENDIX I.B
MANAGEMENT PRACTICE INTERVIEW SCORING GUIDE AND EXAMPLE RESPONSES FOR 4 OF THE 18 PRACTICES

Practice 3: Process problem documentation (operations)

	Score 1	Score 3	Score 5
Scoring grid:	No process improvements are made when problems occur.	Improvements are made in weekly workshops involving all staff, to improve performance in their area of the plant	Exposing problems in a structured way is integral to individuals' responsibilities, and resolution occurs as a part of normal business processes rather than by extraordinary effort/teams
Examples:	A U.S. firm has no formal or informal mechanism in place for either process documentation or improvement. The manager admitted that production takes place in an environment where nothing has been done to encourage or support process innovation.	A U.S. firm takes suggestions via an anonymous box; they then review these each week in their section meeting and decide on any that they would like to proceed with.	The employees of a German firm constantly analyze the production process as part of their normal duty. They film critical production steps to analyze areas more thoroughly. Every problem is registered in a special database that monitors critical processes, and each issue must be reviewed and signed off by a manager.

APPENDIX I.B
(CONTINUED)

	Score 1	Score 3	Score 5
Practice 4: Performance tracking (monitoring)			
Scoring grid:	Measures tracked do not indicate directly if overall business objectives are being met. Tracking is an ad hoc process (certain processes aren't tracked at all)	Most key performance indicators are tracked formally. Tracking is overseen by senior management.	Performance is continuously tracked and communicated, both formally and informally, to all staff using a range of visual management tools.
Examples:	A manager of a U.S. firm tracks a range of measures when he does not think that output is sufficient. He last requested these reports about 8 months ago and had them printed for a week until output increased again. Then he stopped and has not requested anything since.	At a U.S. firm every product is bar-coded and performance indicators are tracked throughout the production process; however, this information is not communicated to workers.	A U.S. firm has screens in view of every line. These screens are used to display progress to daily target and other performance indicators. The manager meets with the shop floor every morning to discuss the day past and the one ahead and uses monthly company meetings to present a larger view of the goals to date and strategic direction of the business to employees. He even stamps canteen napkins with key performance achievements to ensure everyone is aware of a target that has been hit.

APPENDIX I.B
(CONTINUED)

Practice 11: Targets are stretching (targets)

Score 1

Scoring grid: Goals are either too easy or impossible to achieve; managers provide low estimates to ensure easy goals.

Examples: A French firm uses easy targets to improve staff morale and encourage people. They find it difficult to set harder goals because people just give up and managers refuse to work people harder.

Score 3

In most areas, top management pushes for aggressive goals based on solid economic rationale. There are a few “sacred cows” that are not held to the same rigorous standard.

A chemicals firm has two divisions, producing special chemicals for very different markets (military, civil). Easier levels of targets are requested from the founding and more prestigious military division.

Score 5

Goals are genuinely demanding for all divisions. They are grounded in solid economic rationale.

A manager of a UK firm insisted that he has to set aggressive and demanding goals for everyone—even security. If they hit all their targets he worries that he has not stretched them enough. Each KPI is linked to the overall business plan.

APPENDIX I.B
(CONTINUED)

	Score 1	Score 3	Score 5
Practice 16: Promoting high performers (incentives)			
Scoring grid:	People are promoted primarily upon the basis of tenure.	People are promoted upon the basis of performance.	We actively identify, develop, and promote our top performers.
Examples:	A UK firm promotes based on an individual's commitment to the company measured by experience. Hence, almost all employees move up the firm in lock step. Management was afraid to change this process because it would create bad feeling among the older employees, who were resistant to change.	A U.S. firm has no formal training program. People learn on the job and are promoted based on their performance on the job.	At a UK firm each employee is given a red light (not performing), an amber light (doing well and meeting targets), a green light (consistently meeting targets, very high performer), or a blue light (high performer capable of promotion by up to two levels). Each manager is assessed every quarter based on his succession plans and development plans for individuals.

Notes: Any score from 1 to 5 can be given, but the scoring guide and examples are only provided for scores of 1, 3, and 5. Multiple questions are used for each dimension to improve scoring accuracy. The full set of scoring and examples can be found in Bloom and Van Reenen (2006). The survey software is available on <http://www.stanford.edu/~n.bloom/>

APPENDIX I.C
PRACTICE LEVEL AVERAGES BY COUNTRY

Practice name	Practice number	Average value by country (U.S. = 100)			Regression coefficients
		(1)	(2)	(3)	(4)
		UK	Germany	France	All
Modern manufacturing, introduction	1	90.2 (3.50)	86.2 (3.46)	101.1 (3.63)	0.013 (0.011)
Modern manufacturing, rationale	2	93.1 (3.35)	101.4 (3.31)	101 (3.47)	0.012 (0.011)
Process documentation	3	89.2 (3.51)	106.8 (3.47)	99.4 (3.64)	0.044*** (0.012)
Performance tracking	4	98.4 (3.2)	109.1 (3.16)	110.7 (3.32)	0.020* (0.012)
Performance review	5	94.6 (2.99)	109.7 (2.96)	104.3 (3.11)	0.020* (0.012)
Performance dialogue	6	92.8 (3.18)	103 (3.09)	99.2 (3.26)	0.024** (0.013)
Consequence management	7	96.4 (3.02)	108.6 (3)	93.5 (3.14)	0.027** (0.013)
Target breadth	8	91.4 (3.53)	93 (3.49)	94 (3.66)	0.020* (0.012)
Target interconnection	9	94 (3.55)	97.3 (3.51)	77.8 (3.68)	0.035** (0.014)
Target time horizon	10	92.1 (3.7)	98.9 (3.66)	91.8 (3.84)	0.032*** (0.012)
Targets are stretching	11	88.1 (3.33)	104.8 (3.3)	101.4 (3.45)	0.022* (0.012)
Performance clarity and comparability	12	93.4 (3.53)	80.5 (3.48)	82.7 (3.65)	0.008 (0.012)
Managing human capital	13	89.3 (3.94)	98.5 (3.91)	89 (4.09)	0.043** (0.013)
Rewarding high performance	14	81.7 (3.41)	84.9 (3.39)	85 (3.54)	0.024** (0.011)
Removing poor performers	15	89.5 (3.04)	92.5 (3.01)	83 (3.15)	0.016 (0.012)
Promoting high performers	16	90.3 (2.86)	104.7 (2.83)	92.1 (2.97)	0.031** (0.012)
Attracting human capital	17	90.5 (2.9)	94.8 (2.87)	85.2 (3)	0.032*** (0.012)
Retaining human capital	18	93.5 (2.75)	97.5 (2.72)	96.6 (2.84)	0.004 (0.012)
Unweighted average		91.7	98.6	93.9	0.023

Notes. In columns (1) to (3) each practice is benchmarked against the average U.S. score (U.S. = 100). In columns (1) to (3) the standard deviation of each practice's average response is reported below in brackets. Calculated from full sample of 732 firms. Management z-scores used in these calculations. In column (4) results are given from 18 OLS estimations following exactly the same specification as column (4), Table I, except estimated with each individual practice z-score one by one rather than the average management z-score. So every cell in column (4) is from a different regression, where standard errors in parentheses allow for arbitrary heteroscedasticity and correlation (clustered by firm), and regression includes "general controls" as detailed in Table I. *** denotes that the variable is significant at the 1% level, ** denotes 5% significance, and * denotes 10% significance.

APPENDIX II: DATA

The entire anonymized dataset, with a full set of do-files generating all results in this paper, is available online at www-econ.stanford.edu/faculty/bloom.html and <http://cep.lse.ac.uk/people/bio.asp?id=1358>.

II.A. Sampling Frame Construction

Our sampling frame was based on the Amadeus dataset for Europe (United Kingdom, France, and Germany) and the Compustat dataset for the United States. These all have information on company accounting data. We chose firms whose principal industries were in manufacturing and that employed (on average between 2000 and 2003) no less than 50 employees and no more than 10,000 employees. We also removed any clients of the consultancy firm we worked with from the sampling frame (33 out of 1,353 firms).

Our sampling frame is reasonably representative of medium-sized manufacturing firms. The European firms in Amadeus include both private and public firms, whereas Compustat includes only publicly listed firms. There is no U.S. database of privately listed firms with information on sales, labor, and capital. Fortunately, a much larger proportion of firms are listed on the stock exchange in the United States than in Europe, so we are able to go far down the size distribution using Compustat. Nevertheless, the U.S. firms in our sample are slightly larger than those of the other countries, so we are always careful to control for size and public listing in the analyses. Furthermore, when estimating production functions, we allow all coefficients to be different on labor, capital, materials, and consolidation status in each country (see notes to Table I).

Another concern is that we condition on firms where we have information on sales, employment, and capital. These items are not compulsory for firms below certain size thresholds, so disclosure is voluntary to some extent for the smaller firms. By design, the firms in our sampling frame (over 50 workers) are past this threshold for voluntary disclosure (the only exception is for capital in Germany).

We achieved a response rate of 54% from the firms that we contacted, a very high success rate given the voluntary nature of participation. Respondents were not significantly more productive

than nonresponders. French firms were slightly less likely to respond than firms in the other three countries, and all respondents were significantly larger than nonrespondents. Apart from these two factors, respondents seemed randomly spread around our sampling frame.

II.B. Firm-Level Data

The collection of the management data and human resource data is described in the text and in Bloom and Van Reenen (2006). Our firm accounting data on sales, employment, capital, profits, shareholder equity, long-term debt, market values (for quoted firms), and wages (where available) came from Amadeus (France, Germany, and the United Kingdom) and Compustat (United States). For other data fields we did the following:

Materials. In France and Germany, these are line items in the accounts. In the United Kingdom, they were constructed by deducting the total wage bill from the cost of goods sold. In the United States, they were constructed following the method in Bresnahan, Brynjolfsson, and Hitt (2002). We start with costs of good sold (COGS) less depreciation (DP) less labor costs (XLR). For firms that do not report labor expenses expenditures, we use average wages and benefits at the four-digit industry level (Bartelsman, Becker, and Gray [2000] until 1996 and then census average production worker annual payroll by four-digit NAICS code) and multiply this by the firm's reported employment level. This constructed measure is highly correlated at the industry level with materials. Obviously there may be problems with this measure of materials (and therefore value added), which is why we check the robustness of the Table I estimates to measures without materials.

Company Shareholdings. This was manually extracted from the Bloomberg online data service for the ten largest shareholders and the ten largest insider shareholders.

Dates of Incorporation (Age). For UK, French, and German companies, these are provided by Amadeus datasets. For the United States, they were obtained from Dunn and Bradstreet.

Family Ownership Data. The ownership data, directors' data, shareholder information, and family generation were collected from company SEC filings (particularly the DEF14a), company databases (Compustat and ICARUS in the United States; Amadeus in the United Kingdom, France, and Germany), company Web sites, *The International Directory of Company Histories*

(St. James Press), and *Moody's Manuals* (Moody's Investor Service). When these data were missing or ambiguous, they were supplemented with information from the family firm telephone survey, which was run on around 300 firms in the sample that were (or potentially were) family-owned. This allowed us to separate firms into the three family firm categories: "family largest shareholder," "family largest shareholder and family CEO," and "family largest shareholder, family CEO, and primogeniture." "Family largest shareholder" firms were defined as those with a single family (combined across all family members, who are all second generation or beyond) as the largest shareholder; "family largest shareholder and family CEO" firms are defined as the subset of "family largest shareholder" firms where the CEO was a family member. Finally, "family largest shareholder, family CEO, and primogeniture" were the subset of "family largest shareholder and family CEO" firms where the eldest male child was selected as the CEO upon succession. In the regressions of Table V, column (5), the omitted baseline category includes founder firms (114 observations), institutionally owned firms (including banks, pension funds, insurance companies, and private equity—336 observations), manager-owned firms (21 observations), private nonexecutive individually owned firms (59 observations), and others (such as charities, cooperatives and foundations—41 observations). We experiment with dropping the founder firms in column (6) of Table V.

CEO Pay and Age. In the United States, the S&P 1500 largest firms (which cover all sectors) are contained in Execucomp, which provided data for the 106 largest of our U.S. firms. For the remaining firms, we manually downloaded the Def14a proxy statements from the SEC to extract the details of the CEO compensation package and age over the last three accounting years. In the United Kingdom, the highest-paid director is a mandatory line item in the accounts, and we took this as the CEO's salary. In France and Germany we have no data on executive pay.

II.C. Industries and Industry-Level Data

Our basic industry code is the U.S. SIC (1987) three-digit level, which is our common industry definition in all four countries. We allocate each firm to its main three-digit sector (based on sales). For the 732 firms in the sample, we have 105 unique three-digit industries. There are at least two sampled firms in

TABLE A.1
DESCRIPTIVE STATISTICS

	All	France	Germany	UK	U.S.	Number of nonmissing observations
Number of firms, #	732	135	156	151	290	732
Management (mean <i>z</i> -score)	-0.001	-0.084	0.032	-0.150	0.097	732
Employment	2,064	1,065	2,035	1,806	2,526	732
Trade openness (imports/output) in ISIC-2	0.31	0.33	0.32	0.42	0.24	732
Lerner index, excluding the firm itself in three-digit industry	0.055	0.040	0.071	0.040	0.060	726
Number of competitors index, 1 = "none," 2 = "a few," 3 = "many"	2.47	2.32	2.35	2.56	2.56	686
Age of firm (in years)	53.9	38.6	86.8	44.3	48.4	732
Listed firm, %	57.2	16.3	41.0	28.5	100	732
Share of workforce with degrees, %	21.2	15.5	14.3	14.0	31.0	526
Share of workforce with an MBA, %	1.36	0.23	0.09	1.28	2.74	510
Average hours per week, all employees	40.7	35.6	38.6	40.8	44.1	555
Tobin's <i>Q</i> (in 2002)	1.92	1.15	1.88	1.87	2.01	369

each industry for 97.4% of the sample. In specifications where we include a full set of three-digit dummies, there are 19 firms that are absorbed away. As a robustness test we reestimated all equations on the sample for which we have at least five firms per industry (this meant dropping 21% of the main sample). The results are very similar to those reported here and are available on request.

The trade data come from the OECD STAN database of industrial production. This is provided at the country ISIC Revision 3 level and is mapped into the U.S. SIC. The measures of competition

we use are “Import Penetration” = $\ln(\text{Import/Production})$ in every country \times industry pair (i.e., 4 countries and 108 industries imply up to 432 cells). We use the average over 1995–1999. “Lerner index of competition” is constructed, as in Aghion et al. (2005), as the mean of $(1 - \text{profit/sales})$ in the entire database (excluding the firm itself) for every country-industry pair (average over 1995–1999 used).

A set of descriptive statistics broken down by country are in Table A.1. We have 732 possible firms with management data, but there are some missing values on a few of the control variables (e.g., percentage MBAs). In these cases we set the value of the control variable equal to zero when it was missing and include an additional dummy variable to indicate this.

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REFERENCES

- Aghion, Philippe, Nicholas Bloom, Richard Blundell, Rachel Griffith, and Peter Howitt, “Competition and Innovation: An Inverted U Relationship,” *Quarterly Journal of Economics*, 120 (2) (2005), 701–728.
- Athey, Susan, and Scott Stern, “An Empirical Framework for Testing Theories about Complementarity in Organizational Design,” NBER Working Paper 6600, 1998.
- Baily, Martin, Charles Hulten, and David Campbell, “Productivity Dynamics in Manufacturing Plants,” *Brookings Papers on Economic Activity*, 1992 (1992), 187–267.
- Bartel, Ann, Casey Ichinowski, and Kathryn Shaw, “How Does Information Technology Really Affect Productivity? Plant-Level Comparisons of Product Innovation, Process Improvement and Worker Skills,” NBER Working Paper 11773, 2005.
- Bartelsman, Eric, Randy Becker, and Wayne Gray, “The NBER Manufacturing Productivity Database,” National Bureau of Economic Research, 2000.
- Bartelsman, Eric, and Phoebus Dhrymes, “Productivity Dynamics: U.S. Manufacturing Plants, 1972–1986,” *Journal of Productivity Analysis*, 9 (1) (1998), 5–34.
- Bartelsman, Eric, and Mark Doms, “Understanding Productivity: Lessons from Longitudinal Microdata,” *Journal of Economic Literature*, 38 (3) (2000), 569–594.
- Bennedsen, Morten, Kasper Nielsen, Francisco Perez-Gonzales, and Daniel Wolfenzon, “Inside the Family Firm: The Role of Families in Succession Decisions and Performance,” *Quarterly Journal of Economics*, 122 (2) (2007), 647–691.
- Berle, Adolph, and Gardiner Means, *The Modern Corporation and Private Property* (New York, NY: MacMillan, 1932).
- Bertrand, Marianne, Simon Johnson, Krislert Samphantharak, and Antoinette Schoar, “Mixing Family with Business: A Study of Thai Business Groups and the Families behind Them,” Mimeo, Chicago GSB, 2005.
- Bertrand, Marianne, and Sendhil Mullainathan, “Do People Mean What They Say? Implications for Subjective Survey Data,” *American Economic Review Papers and Proceedings*, 91 (2) (May 2001), 67–72.

- Bertrand, Marianne, and Antoinette Schoar, "Managing with Style: The Effect of Managers on Firm Policies," *Quarterly Journal of Economics*, 118 (4) (2003), 1169–1208.
- Black, Sandra, and Lisa Lynch, "How to Compete: The Impact of Workplace Practices and Information Technology on Productivity," *Review of Economics and Statistics*, 83 (3) (2001), 434–445.
- Bloom, Nicholas, Tobias Kretschmer, and John Van Reenen, "Work-Life Balance, Management Practices and Productivity," Center for Economic Performance Special Paper No. 16, 2006.
- Bloom, Nicholas, Raffaella Sadun, and John Van Reenen, "Americans do I.T. Better: U.S. Multinationals and the Productivity Miracle," NBER Working Paper No. 13085, 2007.
- Bloom, Nicholas, and John Van Reenen, "Measuring and Explaining Management Practices across Firms and Nations," Center for Economic Performance Discussion Paper No. 716, 2006.
- Blundell, Richard and Steve Bond, "GMM Estimation with Persistent Panel Data: An Application to Production Functions," *Econometric Reviews*, 19 (3) (2000), 321–340.
- Bresnahan, Tim, Eric Brynjolfsson, and Lorin Hitt, "Information Technology, Workplace Organization and the Demand for Skilled Labor: Firm-Level Evidence," *Quarterly Journal of Economics*, 117 (1) (2002), 339–376.
- Burkart, Mike, Fausto Panunzi, and Andrei Shleifer, "Family Firms," *Journal of Finance*, 58 (5) (2003), 2167–2201.
- Caroli, Eve, and John Van Reenen, "Skill-Biased Organizational Change? Evidence from a Panel of British and French Establishments," *Quarterly Journal of Economics*, 116 (4) (2001), 1449–1492.
- Caselli, Francesco, and Nicola Gennaioli, "Dynastic Management," Centre for Economic Performance Discussion Paper No. 668, 2006.
- Chandler, Alfred, *Scale and Scope: The Dynamics of Industrial Capitalism* (Cambridge, MA: Belknap Press, 1994).
- Dunning, John, *American Investment in British Manufacturing Industry* (London: George Allen, 1958).
- Foster, Lucia, John Haltiwanger, and Chad Syverson, "Reallocation, Firm Turnover and Efficiency: Selection on Productivity or Profitability?" NBER Working Paper 11555, 2005.
- Holtz-Eakin, Douglas, David Joulfaian, and Harvey Rosen, "The Carnegie Conjecture: Some Empirical Evidence," *Quarterly Journal of Economics*, 108 (2) (1993), 413–436.
- Hopenhayn, Hugo, "Entry, Exit and Firm Dynamics in Long-Run Equilibrium," *Econometrica*, 60 (5) (1992), 1127–1150.
- Ichniowski, Casey, Kathryn Shaw, and Giovanna Prenushi, "The Effects of Human Resource Management: A Study of Steel Finishing Lines," *American Economic Review*, 87 (3) (1997), 291–313.
- Jovanovic, Boyan (1982), "Selection and the Evolution of Industry," *Econometrica*, 50 (3) (1982), 649–670.
- Landes, David, *The Unbound Prometheus: Technological Change and Industrial Development in Western Europe from 1750 to the Present* (Cambridge, UK: Cambridge University Press, 1969).
- La Porta, Rafael, Florencio Lopez-de-Silanes, and Andrei Shleifer, "Corporate Ownership around the World," *Journal of Finance*, 54 (2) (1999), 471–517.
- Lazear, Edward, "Performance Pay and Productivity," *American Economic Review*, 90 (5) (2000), 1346–1361.
- Lucas, Robert, "On the Size Distribution of Business Firms," *Bell Journal of Economics*, 9(2) (1978), 508–523.
- Manski, Charles, "Measuring Expectations," *Econometrica*, 72 (5) (2004), 1329–1376.
- Melitz, Marc, "The Impact of Trade on Intra-industry Reallocations and Aggregate Productivity Growth," *Econometrica*, 71 (6) (2003), 1695–1725.
- Menchik, Paul, "Primogeniture, Equal Sharing and the U.S. Distribution of Wealth," *Quarterly Journal of Economics*, 94 (2) (1980), 299–316.
- Morck, Randall, Daniel Wolfenzon, and Bernard Yeung, "Corporate Governance, Economics Entrenchment and Growth," *Journal of Economic Literature*, 43 (3) (2005), 655–720.

- Mundlak, Yair, "Empirical Production Function Free of Management Bias," *Journal of Farm Economics*, 43 (1) (1961), 44–56.
- Nicholas, Tom, "Clogs to Clogs in Three Generations? Explaining Entrepreneurial Performance in Britain since 1850," *Journal of Economic History*, 59 (1999), 688–713.
- Nickell, Steve, "Competition and Corporate Performance," *Journal of Political Economy*, 104 (4) (1996), 724–746.
- Novaes, Walter, and Luigi Zingales, "Bureaucracy as a Mechanism to Generate Information," *Rand Journal of Economics*, 35 (2) (2004), 245–259.
- Olley, Steven, and Ariel Pakes, "The Dynamics of Productivity in the Telecommunications Industry," *Econometrica*, 64 (6) (1996), 1263–1297.
- Perez-Gonzalez, Francisco, "Inherited Control and Firm Performance," Mimeo, Columbia University, 2005.
- Proudfoot Consulting, *Missing Millions: How Companies Mismanage Their Most Valuable Resource* (London: Proudfoot Consulting, 2003).
- Raith, Michael, "Competition, Risk and Managerial Incentives," *American Economic Review*, 93 (4) (2003), 1424–1436.
- Schmidt, Klaus, "Managerial Incentives and Product Market Competition," *Review of Economic Studies*, 64 (2) (1997), 191–213.
- Schwarz, Norbert, and Frita Strack, "Reports of Subjective Well-being: Judgmental Processes and their Methodological Implications," in *Well-Being: The Foundations of Hedonic Psychology*, Daniel Kahneman, Ed Diener, and Norbert Schwarz eds. (New York, NY: Russell Sage Foundation, 1999).
- Stewart, Michael, "Union Wage Differentials, Product Market Influences and the Division of Rents," *Economic Journal*, 100 (403) (1990), 1122–1137.
- Syverson, Chad, "Market Structure and Productivity: A Concrete Example," *Journal of Political Economy*, 112 (6) (2004a), 1181–1222.
- , "Product Substitutability and Productivity Dispersion," *Review of Economics and Statistics*, 86 (2) (2004b), 534–550.
- Villalonga, Belen, and Raphael Amit (2006), "How Do Family Ownership, Control and Management Affect Firm Value?" *Journal of Finance*, 80 (2) (2006), 385–417.
- Vives, Xavier, "Innovation and Competitive Pressure," CEPR Discussion Paper DP4369, 2005.